

Adaptive Frasian Inference under Data–Prior Conflict

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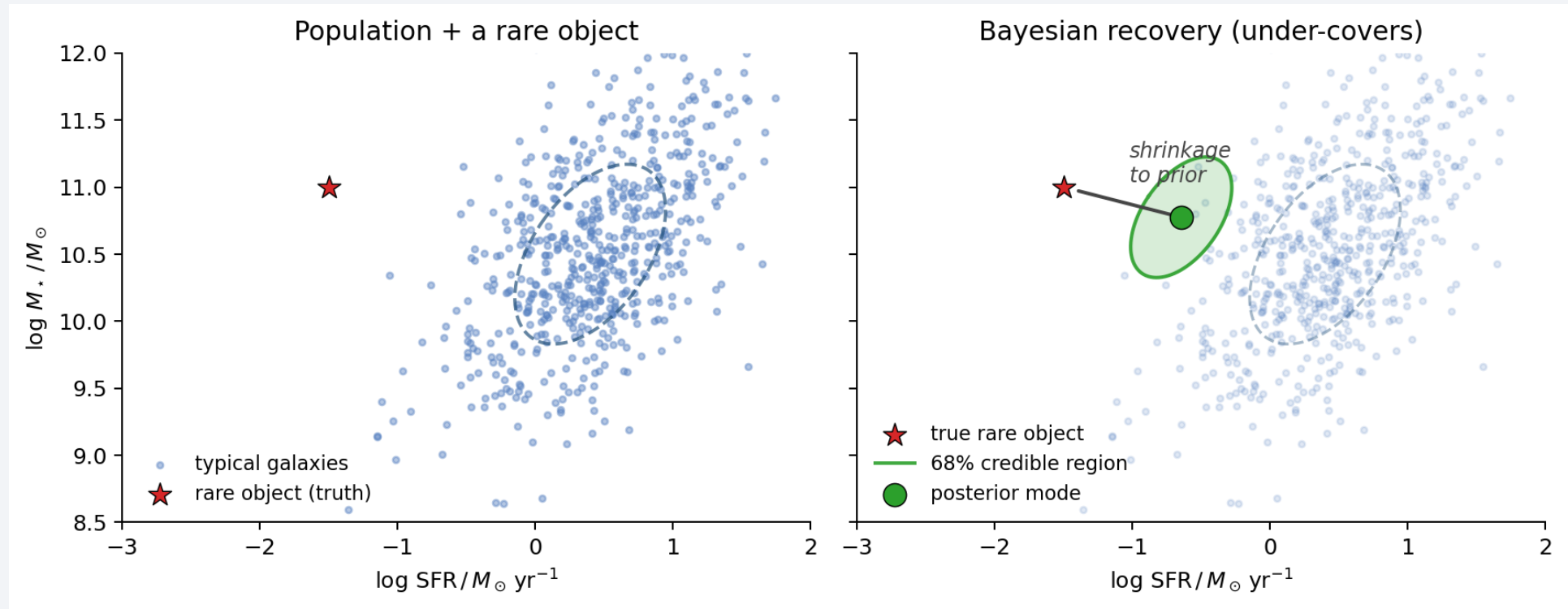
Dunlap Institute for Astronomy & Astrophysics

Data Sciences Institute

University of Toronto

with James Carzon, Ricardo Baptista, & Ann Lee

The interesting galaxies are the ones the prior fights

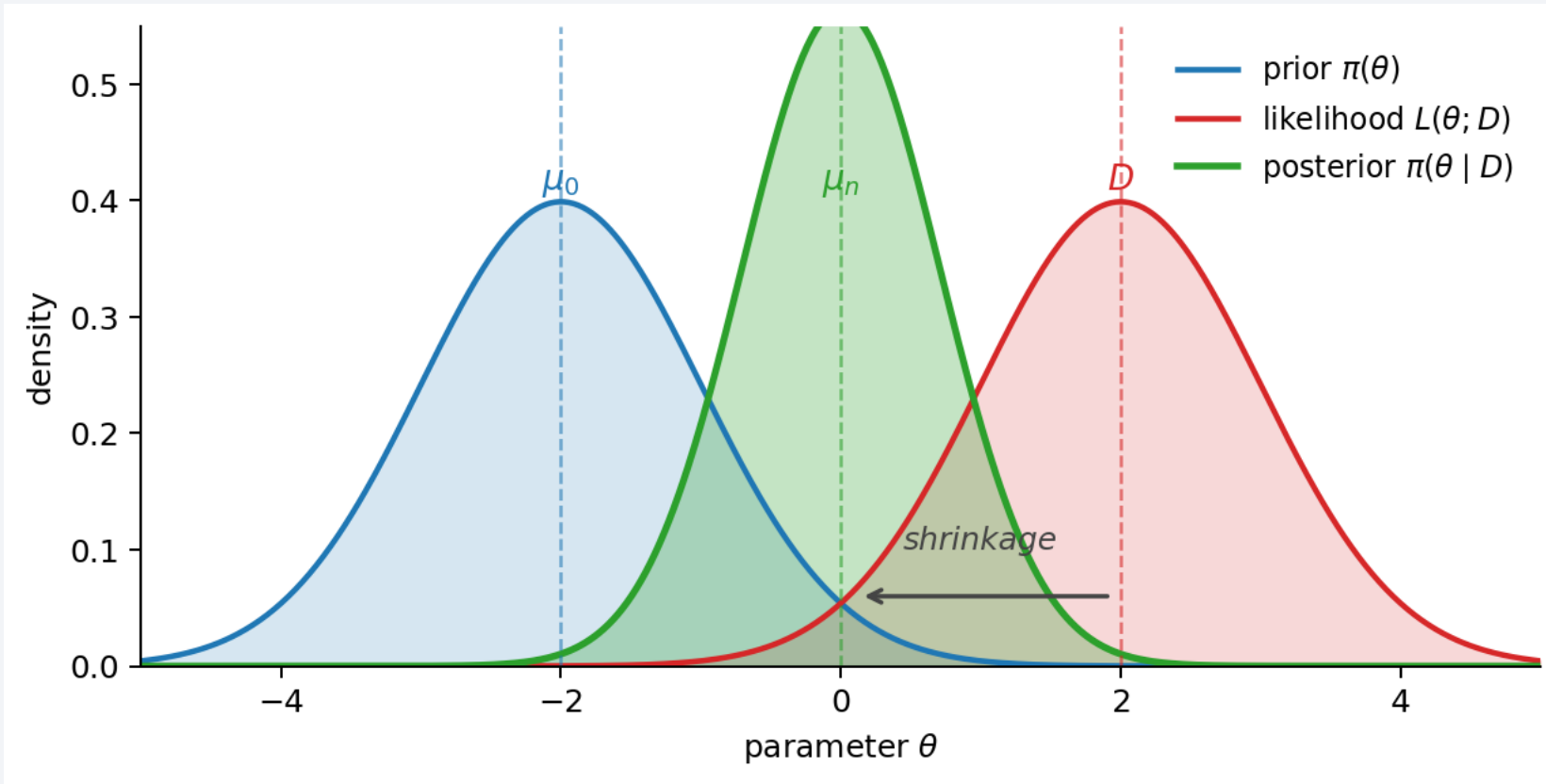


Rare = where the **science** is

Population prior **fights** the data

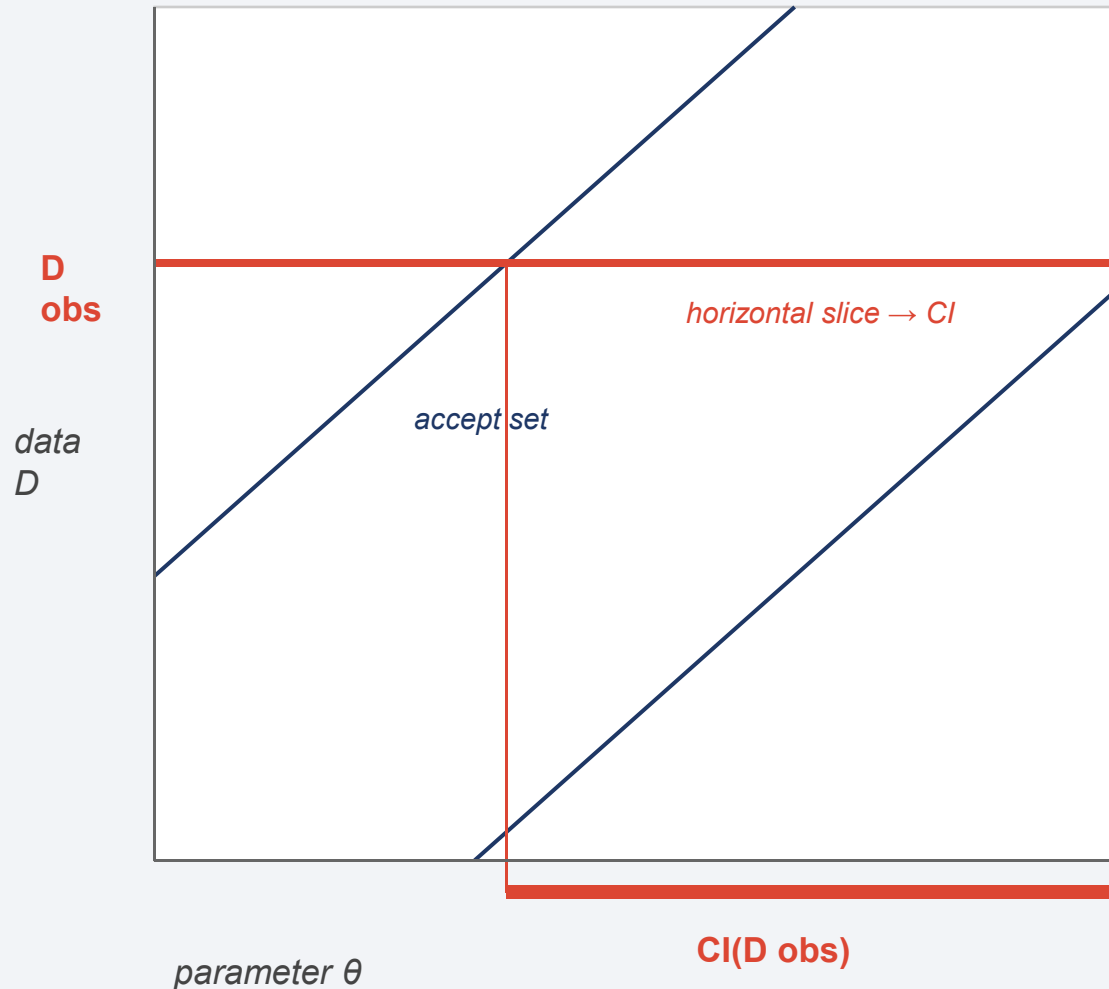
Posterior shrinks to population mean — bias

The prior has information — but neither side uses it well



- **Bayes** lets the prior dominate — biased point
- **Standard frequentist** throws the prior away
- **Want:** use prior when it helps, robust when it doesn't

Test inversion: CIs from p-values

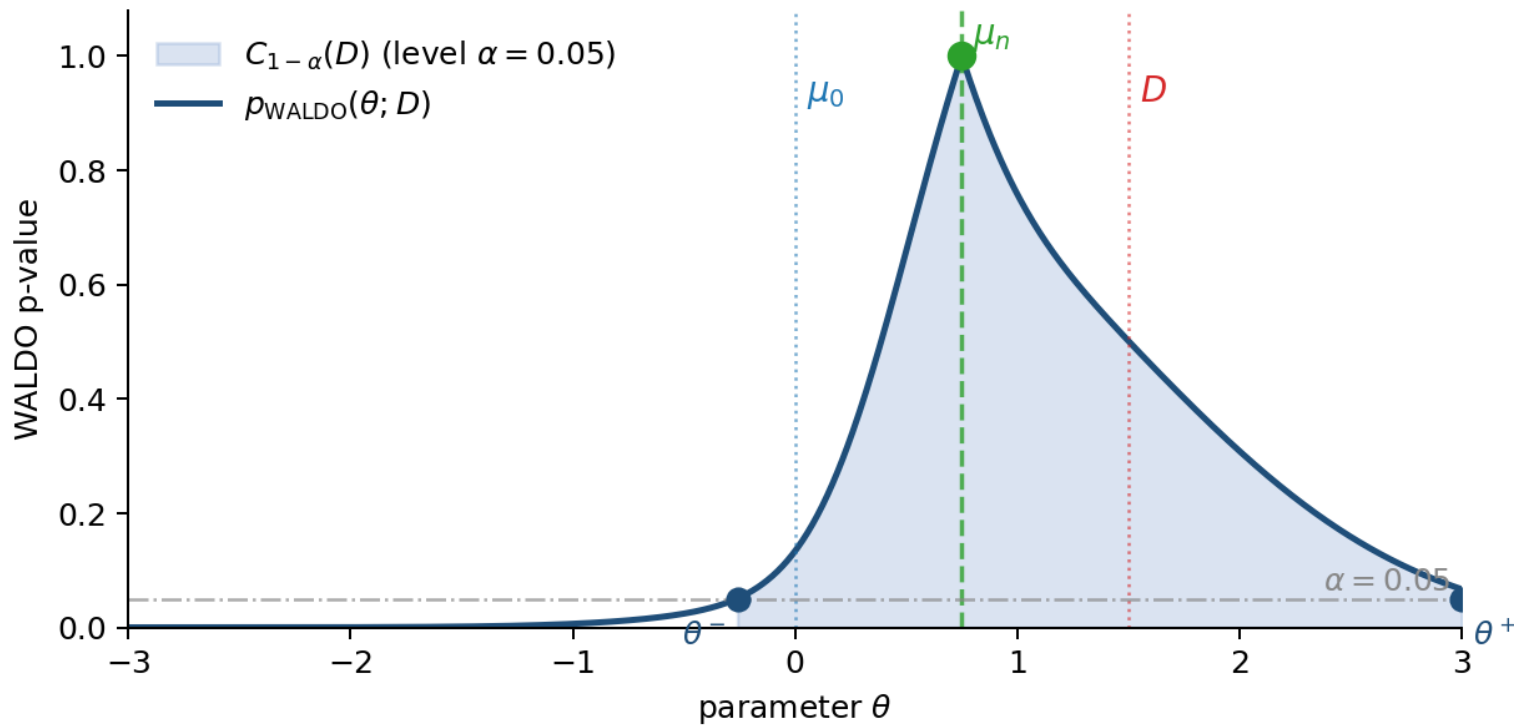


$$CI(D) = \{ \theta : p(\theta; D) \geq \alpha \}$$

- Frequentist machinery, **no Bayesianism required**
- The **test statistic** sets the CI shape
- So: pick a statistic that **uses the prior**

WALDO: a *Frasian* test statistic

WALDO p-value on NN+Normal ($\mu_0 = 0.0, \sigma_0 = 1.0, D = 1.5, \sigma = 1.0$)

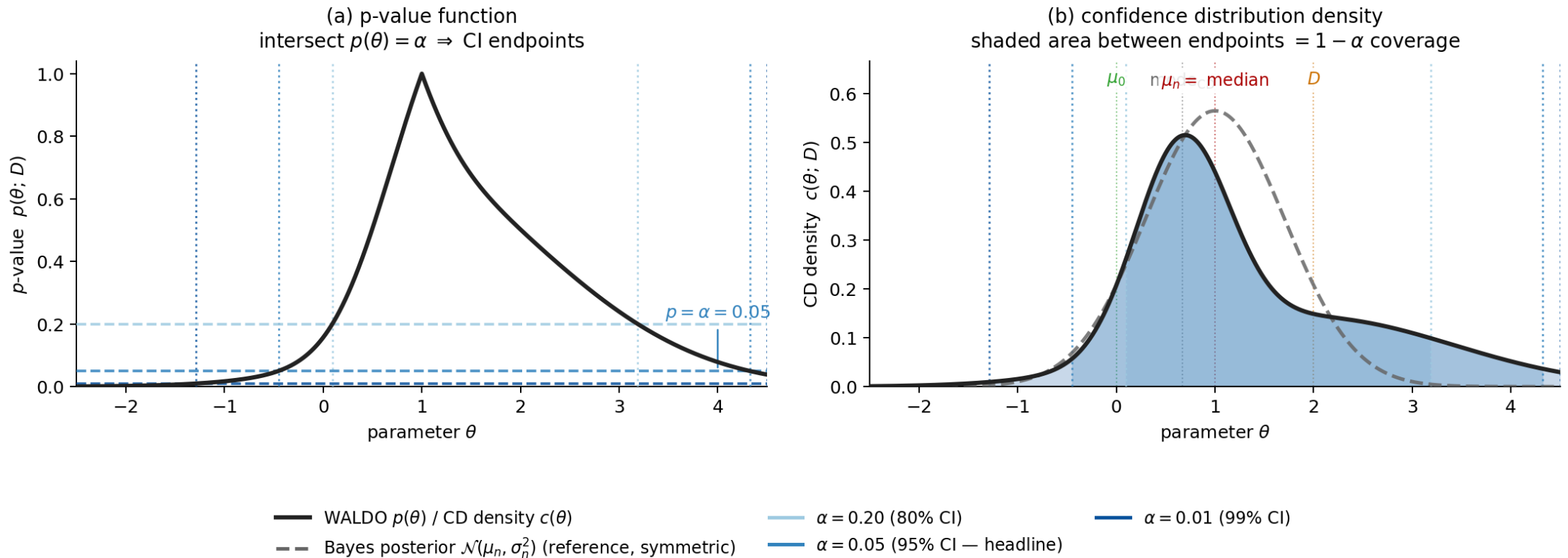


$$T = (\theta - \mu_{\text{post}})^2 / \sigma_{\text{post}}^2$$

- Statistic from the **posterior**
- Calibrated **frequentist**
- *Frasian* = freq. test + Bayes statistic

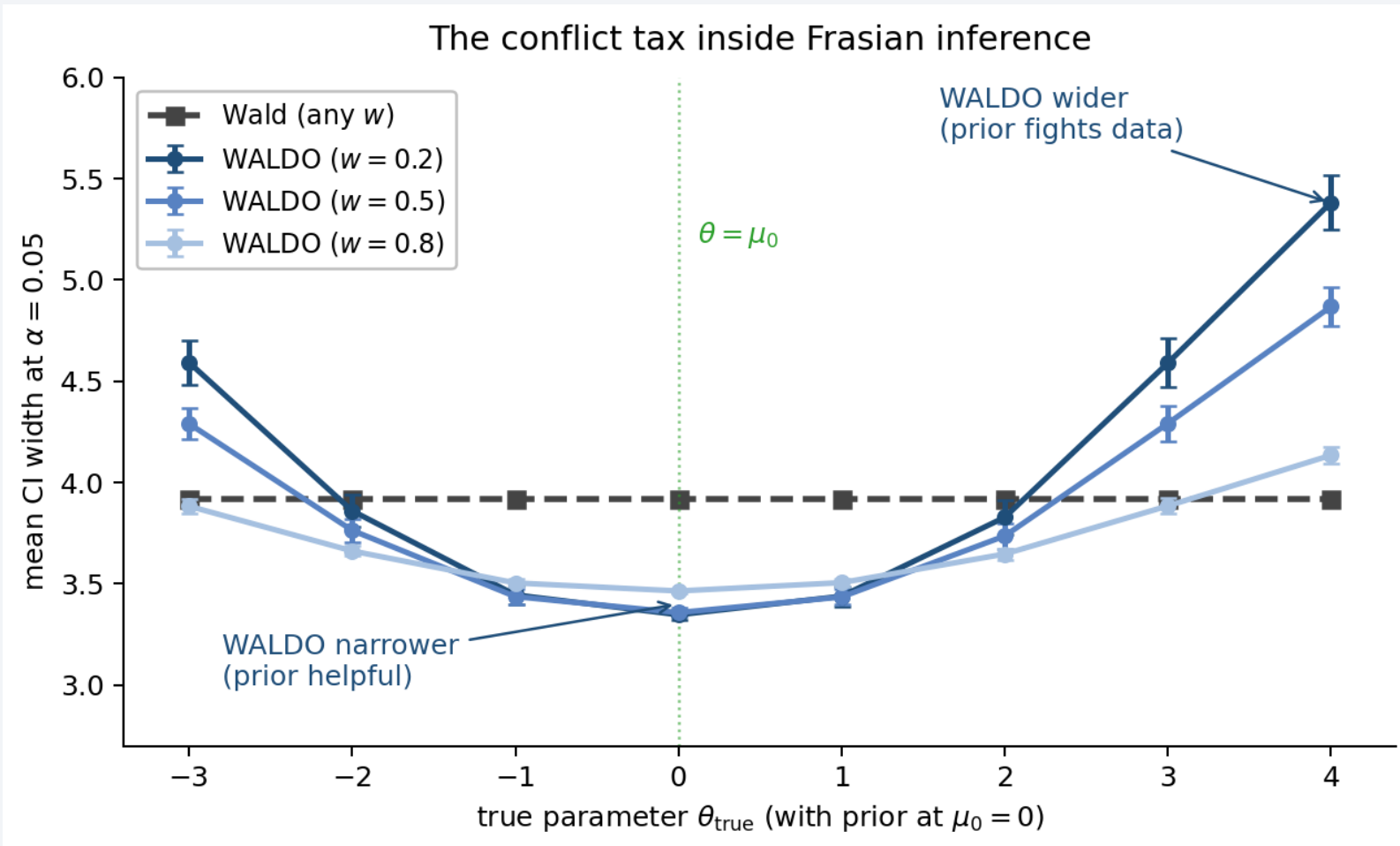
Stacking CIs across α gives a **confidence distribution**

From p-value function to confidence distribution (WALDO on NN+Normal, $D = 2$, $\mu_0 = 0$, $\sigma_0 = \sigma = 1$)



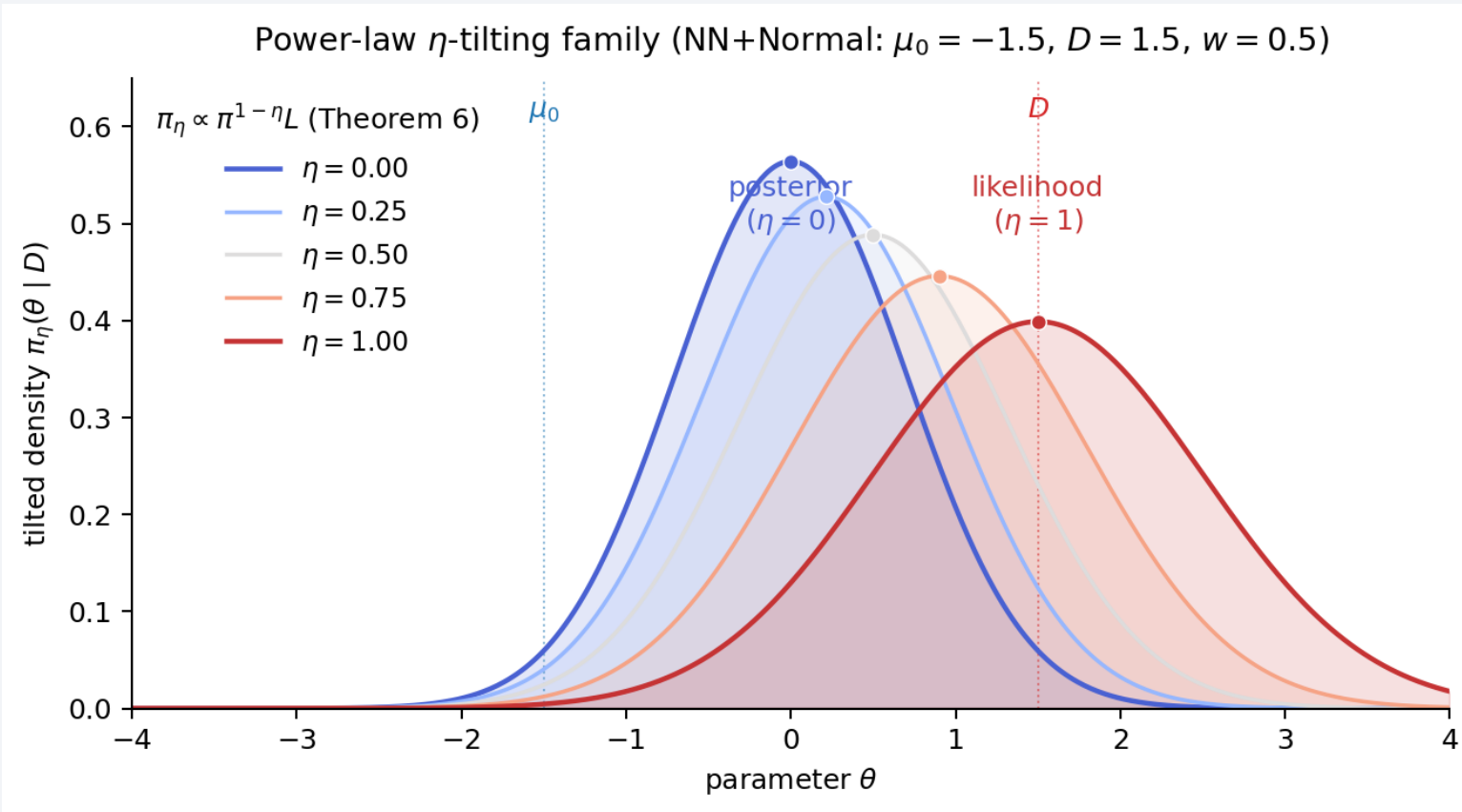
Same coverage guarantee, posterior-shaped density
WALDO CD is asymmetric: mode toward μ_0 , median at μ_n

Coverage holds — but width pays a **conflict tax**



- **Coverage** holds everywhere ✓
- **Width** grows with $|\Delta|$
- **~30%** wider than Wald at moderate conflict
- The ***conflict tax***

A power-law η -knob between WALDO and Wald



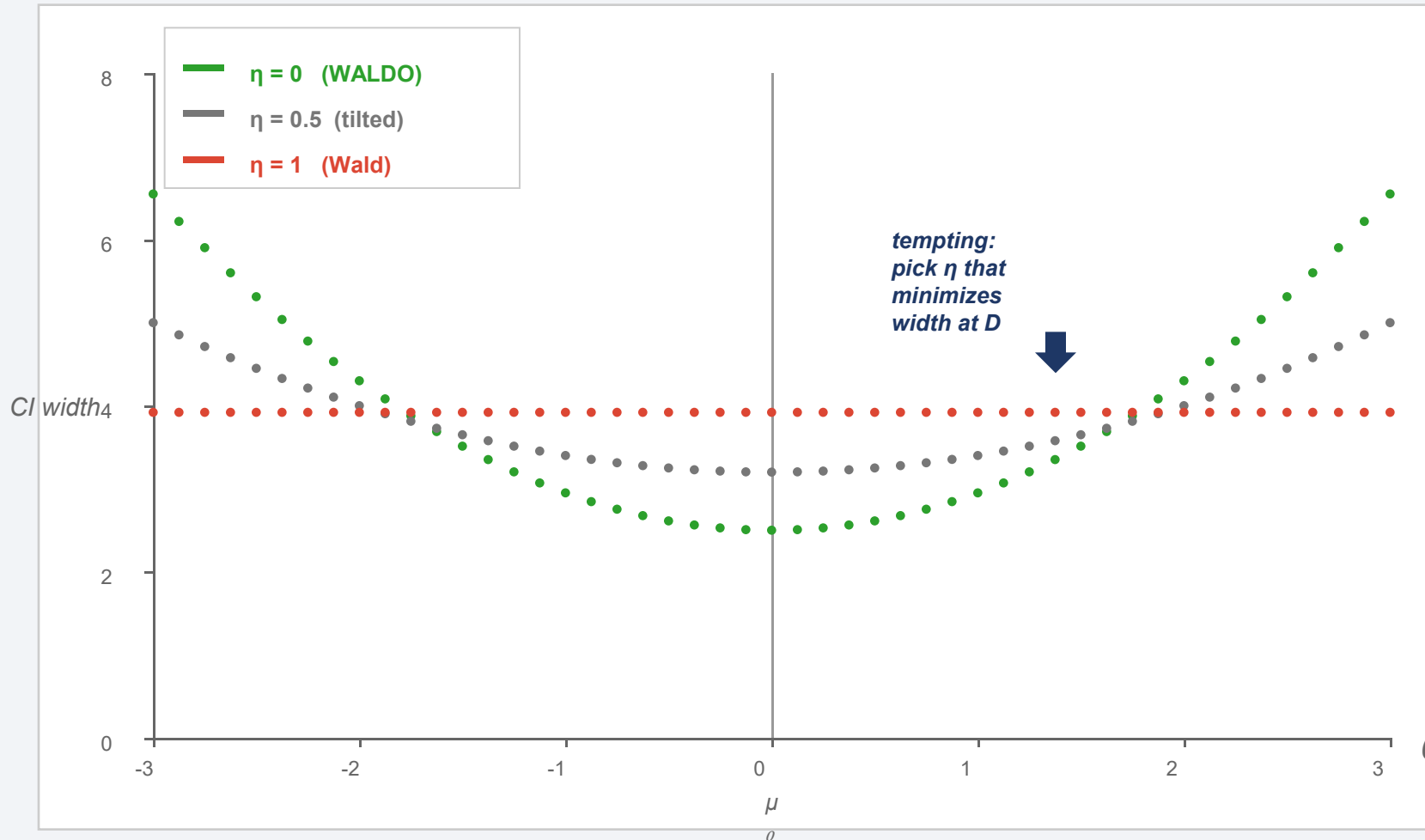
$$\pi_\eta(\theta|D) \propto \pi(\theta)^{1-\eta} \cdot L(\theta;D)$$

- $\eta = 0$: full prior (WALDO)
- $\eta = 1$: no prior (Wald)

How to pick $\eta(\cdot)$ — later

Each η gives a different CI

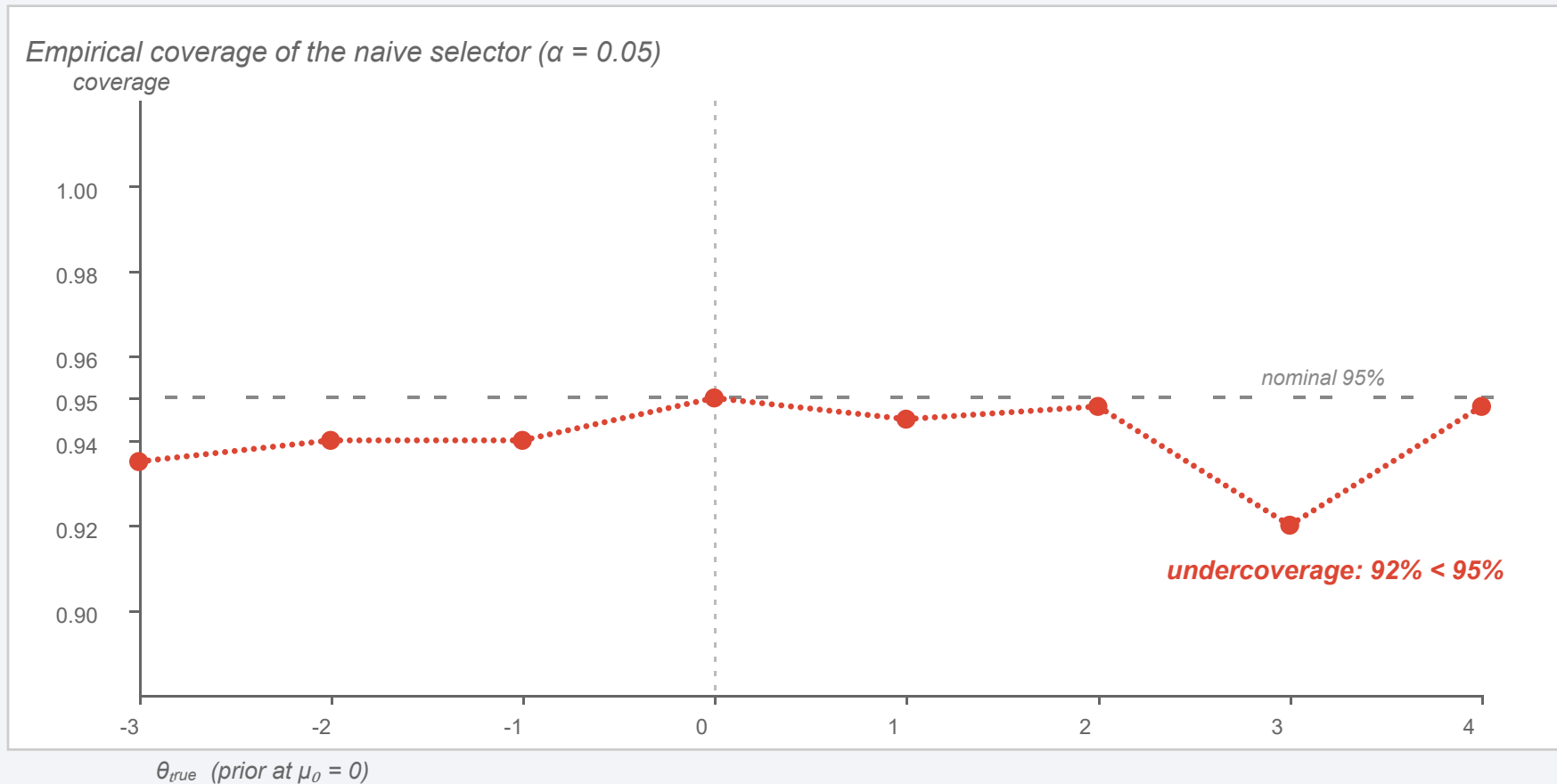
CI width as a function of θ , for fixed η



- Different η \rightarrow different CI width profile
- All are **calibrated** — they trade width for shape
- Tempting: at each D , pick the tightest η
- (spoiler — this breaks calibration)

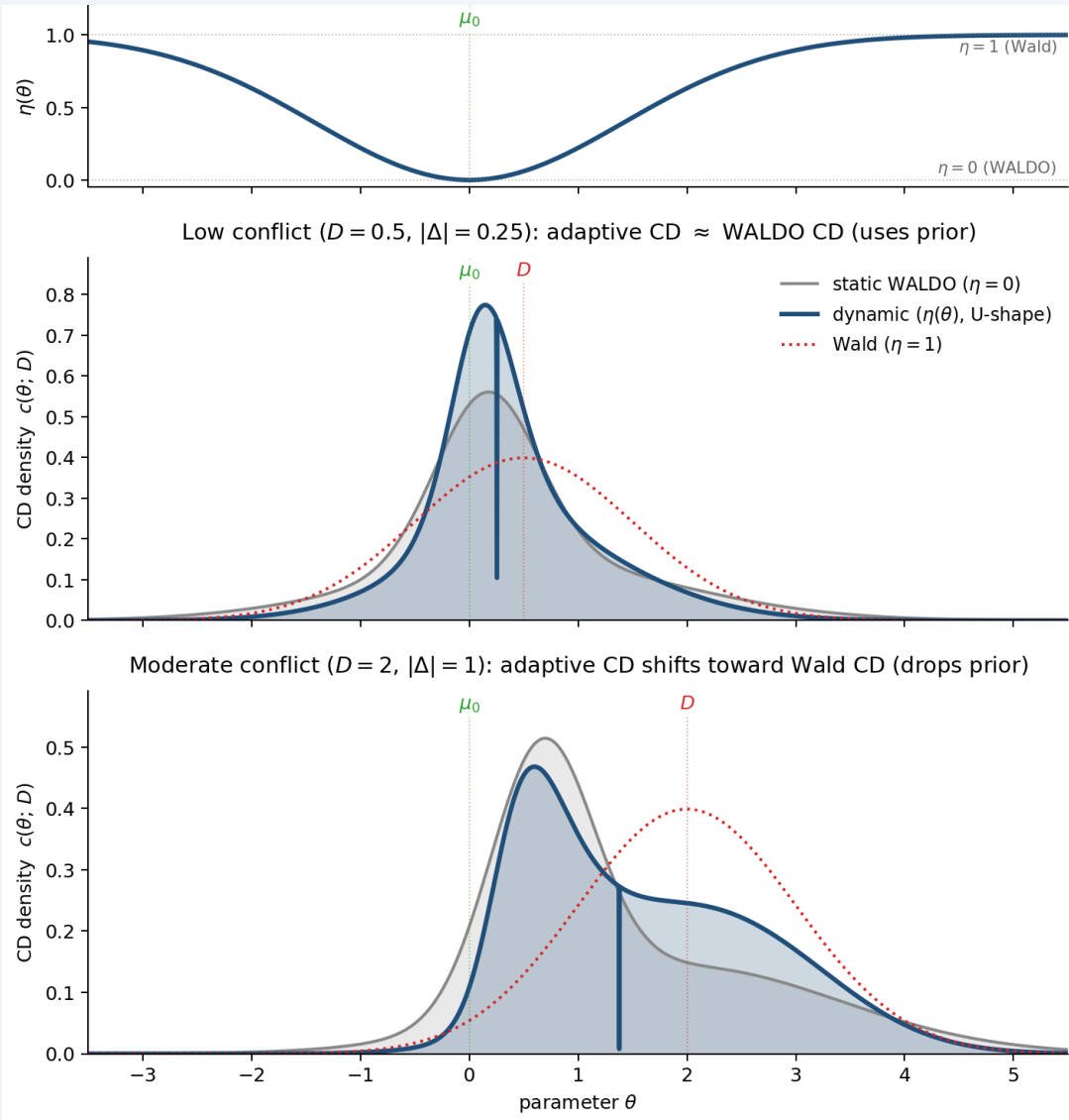
Schematic: CI width for power-law family on NN+Normal at fixed $(\mu_0, \sigma_0, \sigma, w)$.

The naive recipe breaks



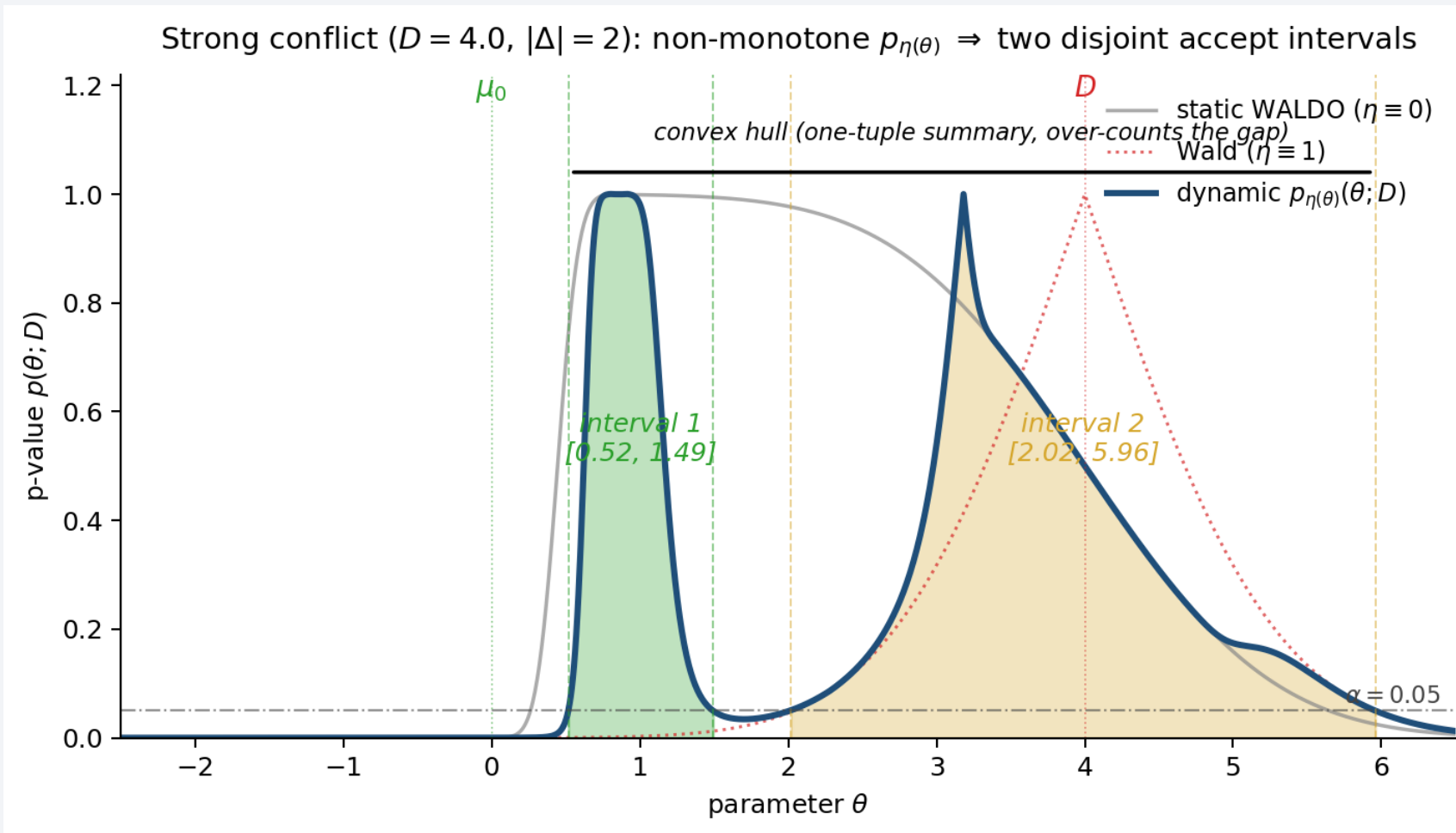
- Recipe: pick $\eta^* = \operatorname{argmin} \operatorname{width}(\eta; D)$
- Empirical coverage: **~93%** — below nominal 95%
- Data picked the test, then tested itself
- **Post-selection inference** — guarantee gone

Dynamic tilting: $\eta(\theta)$, not $\eta(D)$



- **Fix:** let η depend on θ alone — **not D**
- p-value at each fixed η stays $U[0,1]$
✓ calibrated
- Natural shape: U-curve — use prior near μ_0 , drop it at extremes?
- Would allow CD to **adapt** to local prior–data conflict

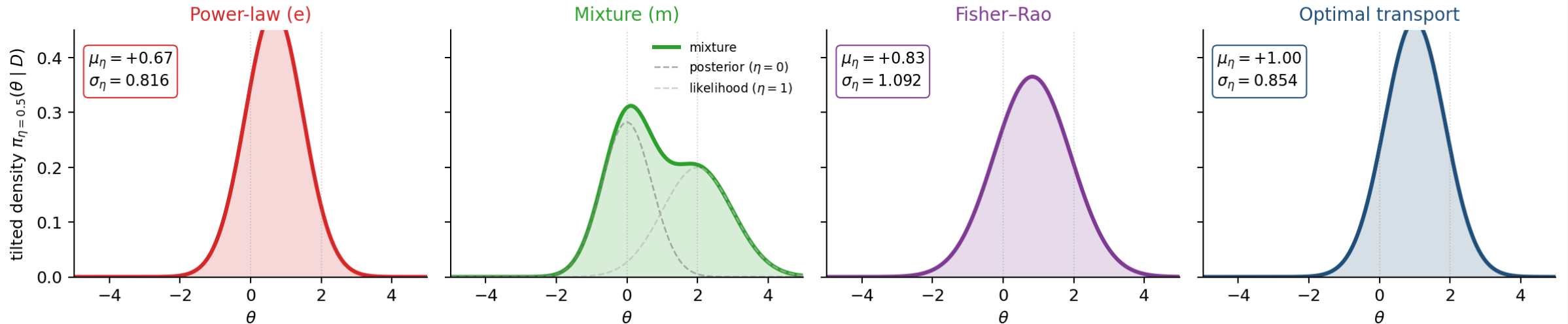
Consequences: disjoint accept regions



- Strong conflict \rightarrow **non-monotone** $p(\theta; D)$
- **Accept region:** union of **two disjoint** intervals
- One near μ_0 , one near D
- Framework returns both — not a convex hull
- **Coverage still exact** — *the truth of the test*

Multiple ways to interpolate between prior and likelihood

Tilted density at $\eta = 0.5$ across schemes (NN+Normal: $\mu_0 = -2.0$, $D = 2.0$, $w = 0.5$)



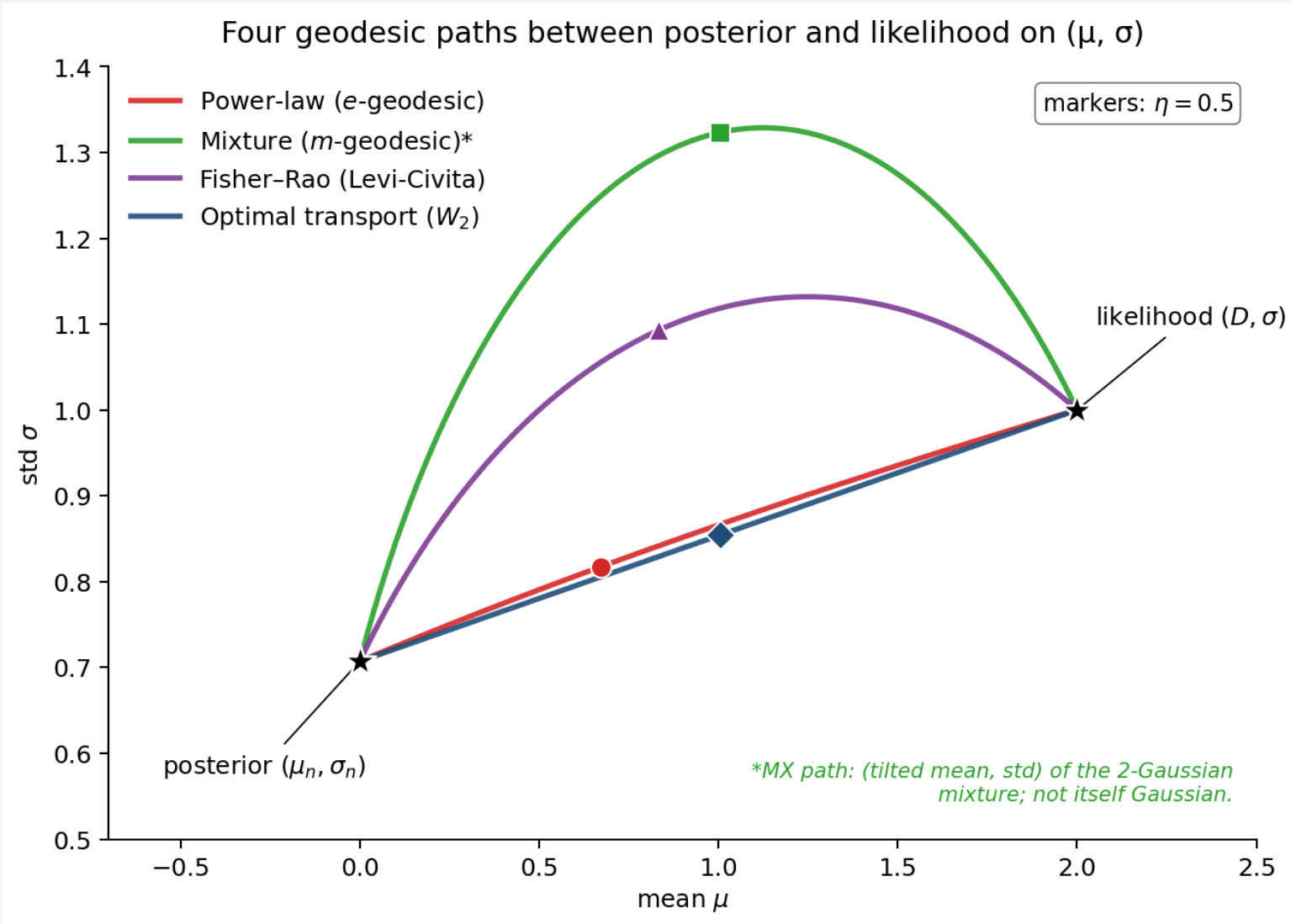
PL $\pi_n \propto \pi^{1-n} \cdot L$
log-linear (e-geodesic)

MX $\pi_n = (1-\eta)\pi_a + \eta L/Z$
linear in densities (m-geodesic)

FR $ds^2 = (d\mu^2 + 2d\sigma^2)/\sigma^2$
Fisher-Rao metric on (μ, σ)

OT $F_n^{-1}(u) = (1-\eta)F_a^{-1} + \eta F_1^{-1}$
linear in quantiles (W_2)

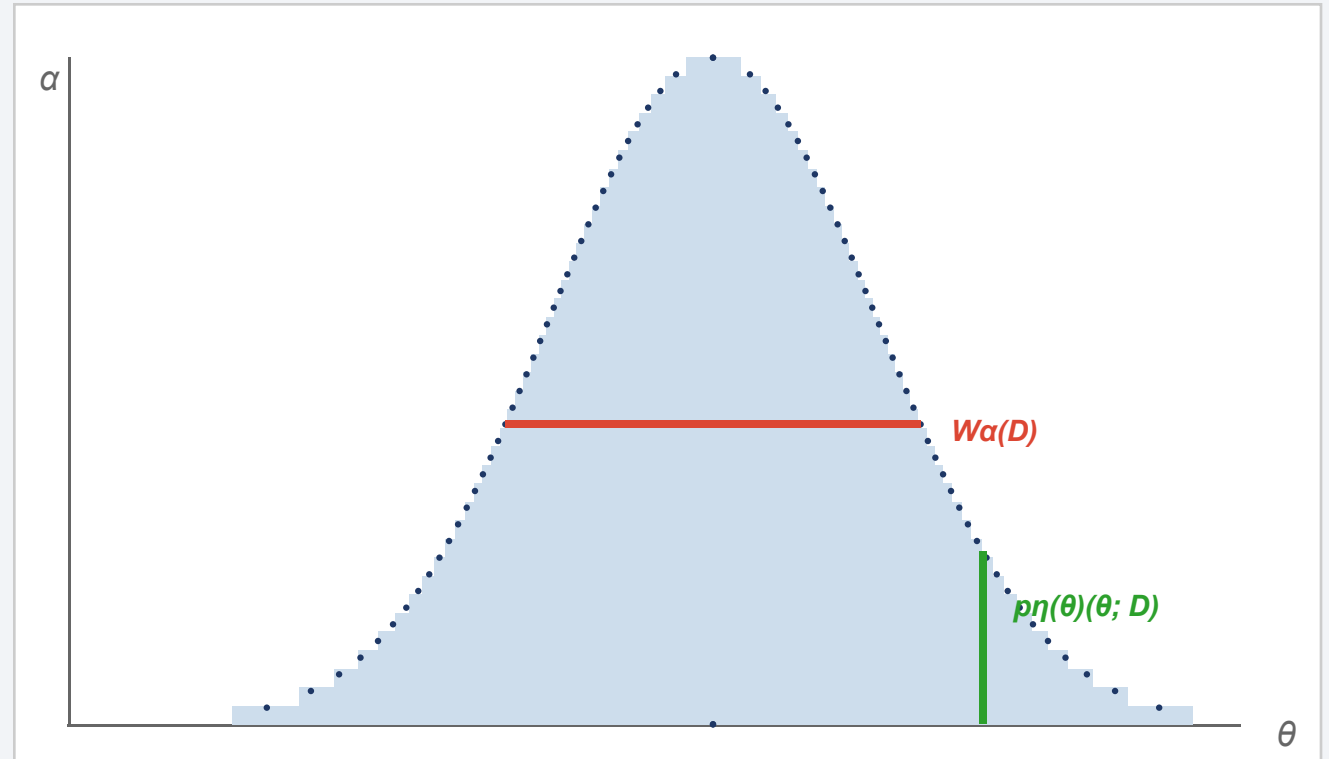
Each family is a path through distribution space



- η traces a **path** from posterior to likelihood
- **PL** \approx **OT** here — Gaussian coincidence
- **MX** bulges up: mixture variance $>$ either endpoint
- **FR** arcs through Fisher–Rao curvature
- Choice = **smoothness** \times **admissibility**, not validity

What objective are we optimizing?

- A good cross- θ functional is **average CI width over a calibration set**
 - Swap order of integration (**Fubini**)
 - \rightarrow trainable as **integrated p-value**



integrate by rows = average width · by columns = integrated p-value

Fubini swap

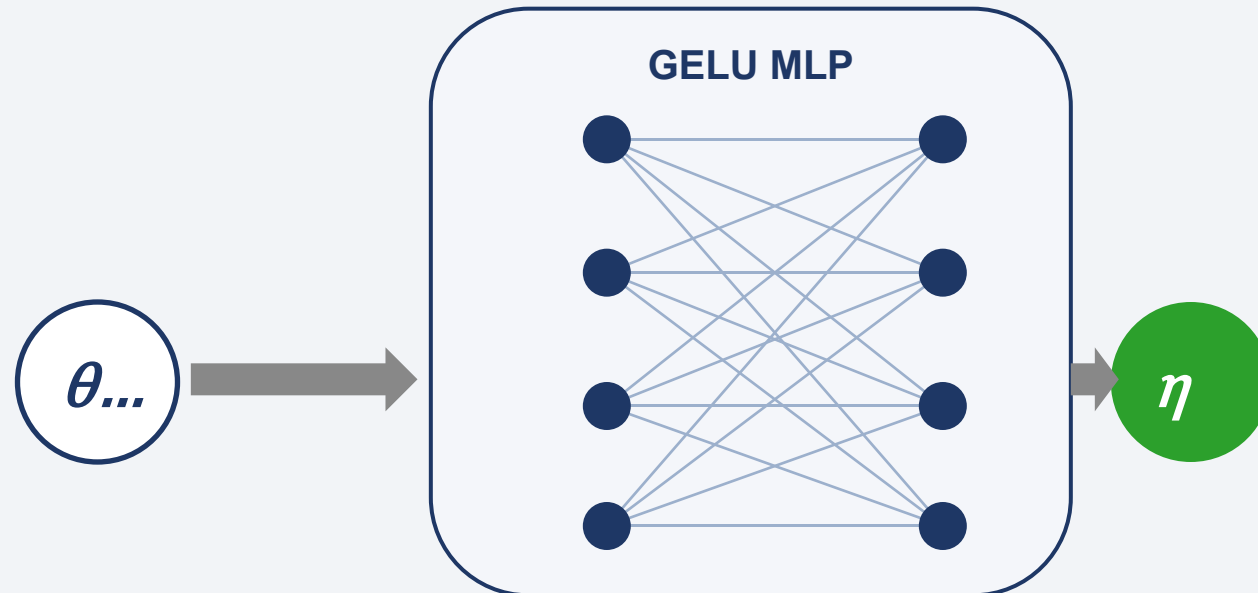
$$\int_{\theta} p_{\eta(\theta)}(\theta; D) d\theta = \int_0^1 W_{\alpha}(D; \eta) d\alpha$$

integrated p-value over θ = average CI width across α -levels

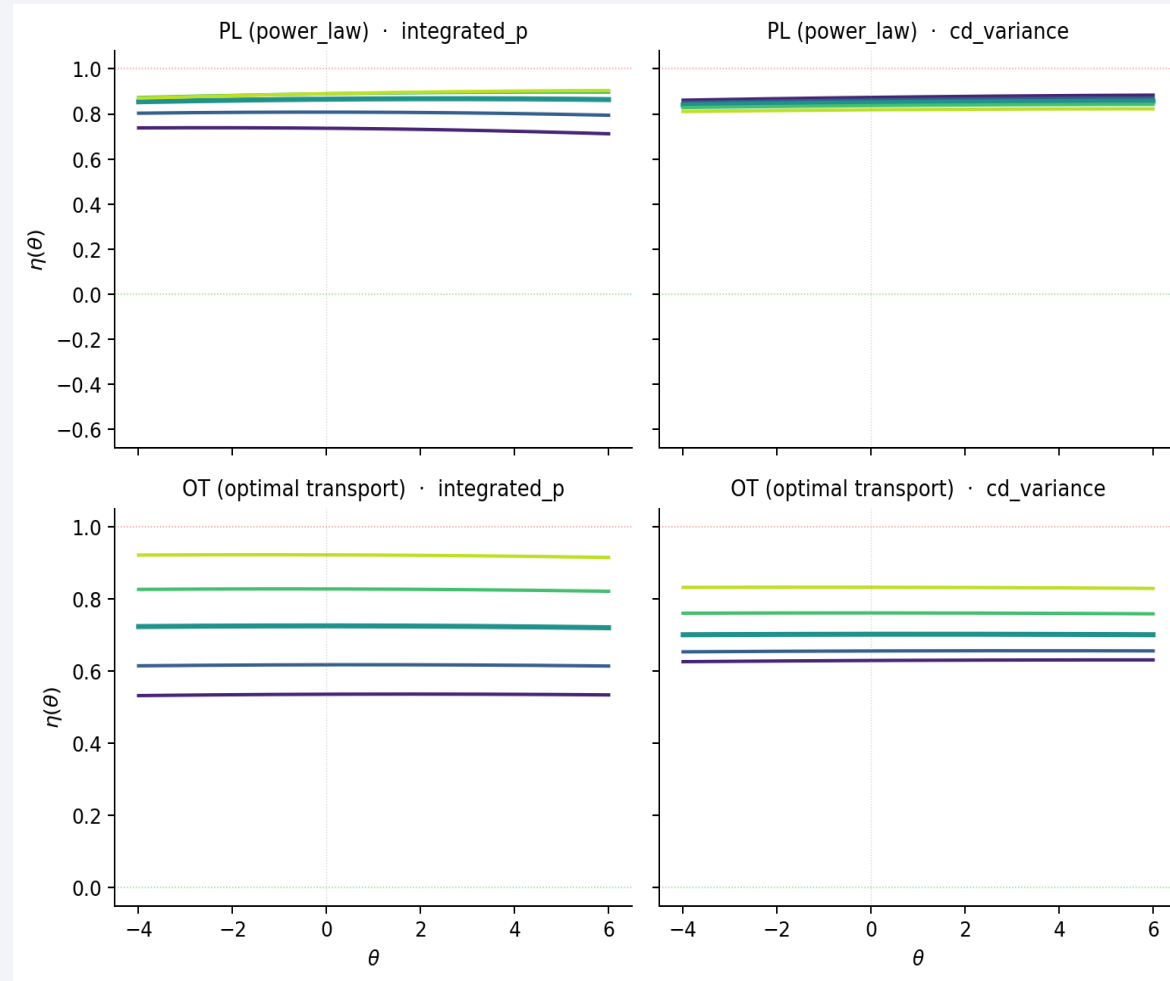
Alternative losses for ablation: **cd_variance** (variance of the CD density) · **static_width** (CI width at a reference θ_0)

Parameterize $\eta(\theta)$ with a small neural net

- $\eta(\cdot)$ is a **function** — needs an approximator
- Small **GELU MLP** mapping $\theta \rightarrow \eta$
- Condition on $(\mu_0, \sigma_0, \sigma)$ — one net per regime
- Second head **ValidityNet**: soft admissibility



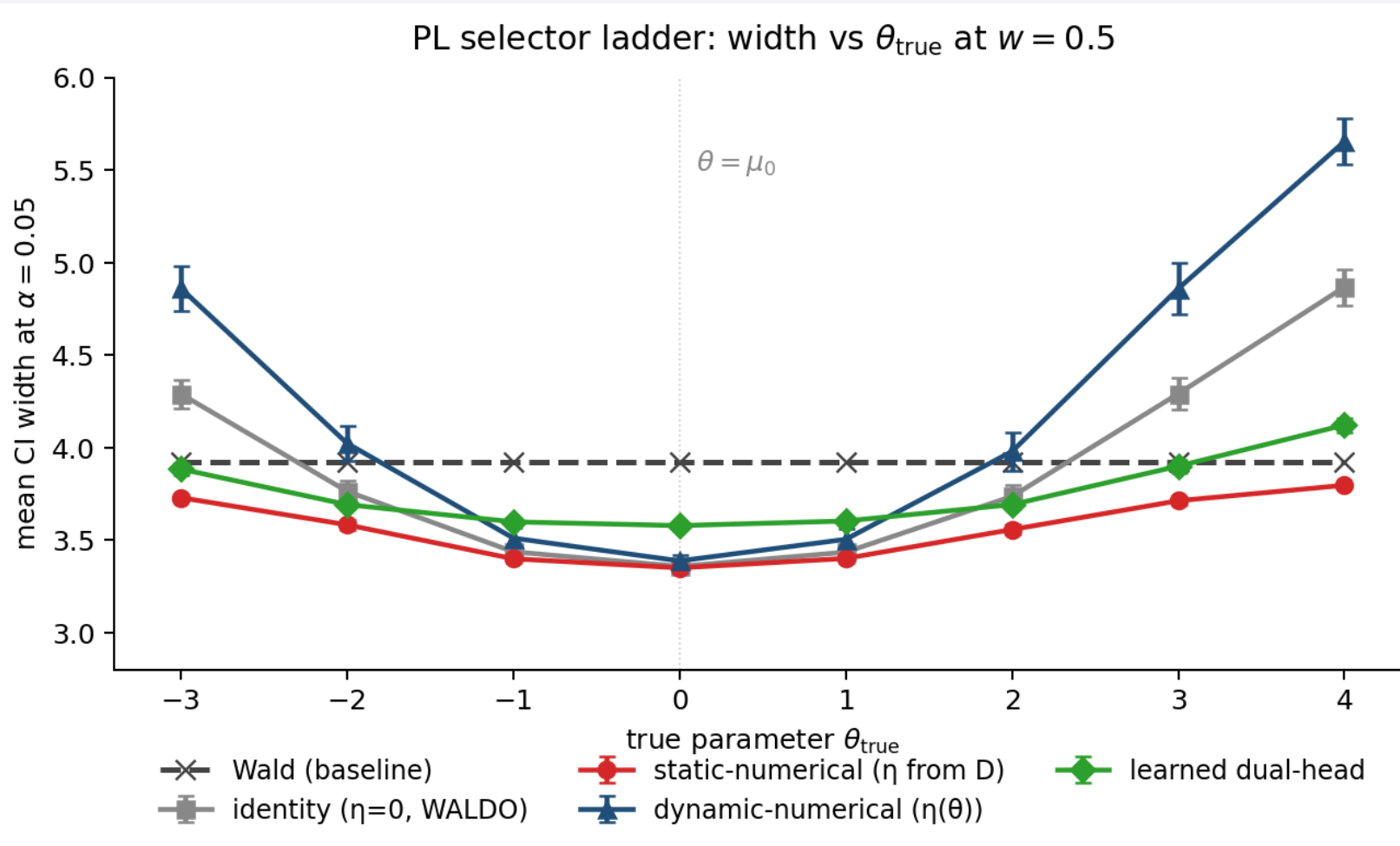
Both schemes adapt — just at different η



Top (PL): η lives near its admissibility cap (~ 0.8) · **Bottom (OT):** η spans 0.4–1.0 across σ_0

Same headline (calibrated, narrower than Wald) — different transit point between posterior and likelihood

Calibrated and narrower than Wald



Coverage ✓ nominal at every θ

Width ✓ 3–9% narrower than Wald at $|\theta| \leq 2$
~5% wider only at $|\theta|=4$ — no post-selection tax

(the static-numerical and dynamic-numerical show what happens in this case)

Only the learned selector hits this target

Summary and outlook

The reframe

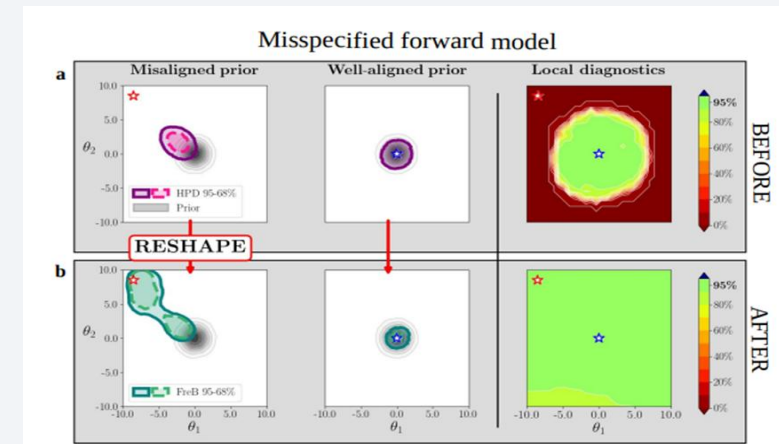
- Bayesian **shape** + frequentist **calibration** — not an either/or
- Treat the prior as **flexible meta-belief** — a dial, not a commitment

What we built

- **η -tilting**: a tunable interpolation between posterior and likelihood
- **$\eta(\theta)$** learned globally — calibration preserved by Neyman inversion
- Result: **calibrated and tighter than Wald** — geometry-agnostic

What's next

- SBI under misspecification — *Carzon et al. 2025, Ingram et al. (in prep.)*
- Multi-dim θ , non-conjugate likelihoods, OOD detection



References / Q&A

WALDO and Frasian lineage

- Masserano, Dorigo, Izbicki, Kuusela, Lee 2023 — *WALDO*
- Dalmaso et al. 2020 — *LF2I*
- Fraser 1968 — *structural inference*
- Dempster 1974, 1997 — *posterior likelihood ratio*
- Aitkin 1991, 1997 — *posterior Bayes factors*
- Wasserman 2011 — *coined "Frasian"*

Post-selection inference

- Berk, Brown, Buja, Zhang, Zhao 2013 — *valid PSI*
- Fithian, Sun, Taylor 2014 — *exact PSI*

Confidence distributions & fusion

- Schweder & Hjort 2016 — *Confidence, Likelihood, Probability*
- Xie & Singh 2013 — *CD review*
- Singh, Xie, Strawderman 2005 — *confidence fusion*

Geometry of distributions

- Amari & Nagaoka 2000 — *Methods of Information Geometry*
- Villani 2003 — *Topics in Optimal Transportation*
- Behboodian 1970 — *Gaussian mixture bimodality*

Prior-data conflict

- Evans & Moshonov 2006; Marshall & Spiegelhalter 2007

Applications & ongoing work

- Johnson, Leja et al. 2021; Leja et al. 2017 — *Prospector*
- Wang et al. 2023 — *Prospector- β*
- Suess et al. 2022; Narayanan et al. 2024 — *atypical galaxies*
- Carzon, Speagle, Lee, Izbicki, Masserano et al. 2025 — *FreB / trustworthy SBI*
- **Ingram et al. (in prep.)** — *selection/calibration*
- Cranmer, Brehmer, Louppe 2020 — *SBI overview*

Foundations

- Casella & Berger 2002 — *Statistical Inference (textbook)*