<u>Title:</u> Affective brain patterns as multivariate neural correlates of cardiovascular disease risk

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Abstract

This study tested whether brain activity patterns evoked by affective stimuli relate to individual differences in an indicator of preclinical atherosclerosis: carotid artery intima-media thickness (CA-IMT). Adults (aged 30-54 years) completed functional magnetic resonance imaging (fMRI) tasks that involved viewing three sets of affective stimuli. Two sets included facial expressions of emotion, and one set included neutral and unpleasant images from the International Affective Picture System (IAPS). Cross-validated, multivariate, and machine-learning models showed that individual differences in CA-IMT were partially predicted by brain activity patterns evoked by unpleasant IAPS images, even after accounting for age, sex, and known cardiovascular disease risk factors. CA-IMT was also predicted by brain activity patterns evoked by angry and fearful faces from one of the two stimulus sets of facial expressions, but this predictive association did not persist after accounting for known cardiovascular risk factors. The reliability (internal consistency) of brain activity patterns evoked by affective stimuli may have constrained their prediction of CA-IMT. Distributed brain activity patterns could comprise affective neural correlates of preclinical atherosclerosis; however, the interpretation of such correlates may depend on their psychometric properties, as well as the influence of other cardiovascular risk factors and specific affective cues.

Key words: affect; cardiovascular disease risk; emotion; fMRI; machine learning

Is there a brain phenotype—a neural pattern expressed reliably by individuals—that relates to risk for atherosclerotic cardiovascular disease (CVD)? Answering this question may not only add to a brain-based and mechanistic understanding of CVD, but may also inform. approaches to predict and reduce vulnerability to a leading cause of population morbidity and mortality (Gianaros & Jennings, 2018; Kraynak, Marsland, & Gianaros, 2018). Existing evidence suggests that CVD risk relates to individual differences in the functionality of cortical and limbic brain systems that encode, process, and orchestrate responses to affective cues and contexts (Kraynak, Marsland, & Gianaros, 2018). Through their interplay with brainstem cell groups, these cortical and limbic systems are specifically postulated to link affective information processing to downstream control over autonomic, immune, and neuroendocrine parameters of peripheral physiology that may be pathogenic to the heart and blood vessels (Kraynak, Marsland, & Gianaros, 2018). Accordingly, quantifying reliable individual differences in the functionality of brain systems for affective processing may help to identify novel neural correlates of CVD risk. Such affective neural correlates could plausibly complement the use of conventional risk factors in the prediction of CVD. They may also help to provide a brain-based account of epidemiological associations between CVD and peripheral physiological and behavioral correlates of mood and emotion (e.g., Davidson, Alcantara, & Miller, 2018; DeSteno, Gross, & Kubzansky, 2013; Krantz & McCeney, 2002; Rozanski, Blumenthal, & Kaplan, 1999; Suls & Bunde, 2005).

At present, several autonomic, immune, and neuroendocrine correlates of CVD risk have been associated with the functionality of a particular brain system—the amygdala (for reviews, see Gianaros & Wager, 2015; Ginty, Kraynak, Fisher, & Gianaros, 2017; Kraynak, Marsland, Wager, & Gianaros, 2018; Muscatell & Eisenberger, 2012; Ruiz Vargas, Sörös, Shoemaker, & Hachinski, 2016; Thayer, Ahs, Fredrikson, Sollers, & Wager, 2012). Moreover, individual

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differences in amygdala function have been associated with clinical cardiovascular endpoints, cardiometabolic outcomes that are conducive to atherosclerotic CVD, and even surrogate markers of preclinical atherosclerosis that are all presaged by autonomic, immune, and neuroendocrine dysregulation. Elevated resting metabolic activity of the amygdala, for example, confers risk for cardiovascular events (Tawakol et al., 2017); incident diabetes (Osborne et al., 2019); the progression of visceral adiposity (Ishai et al., 2019); and, non-calcified coronary plaque (Goyal et al., 2018). The latter findings build on earlier cross-sectional evidence that preclinical atherosclerosis in the carotid arteries associates with elevated metabolic activity in the medial temporal lobe (Sojkova et al., 2010) and – more specifically – with amygdala reactivity to facial expressions of fear and anger (Gianaros et al., 2008). Elevated amygdala activity at rest or evoked by affective cues may thus comprise a neural correlate of CVD risk.

Not all findings, however, support the latter possibility. Hence, the severity of preclinical atherosclerosis was not statistically associated with extracted and averaged levels of amygdala reactivity to unpleasant affective scenes (Gianaros et al., 2014). Furthermore, no statistical associations were observed between conventional indicators of CVD risk (e.g., Framingham risk scores, metabolic syndrome) and stressor-evoked activity in cortical and limbic brain systems, including the amygdala (Lederbogen et al., 2018). Consequently, a still open question is whether CVD risk relates to a reproducible neural correlate that encompasses amygdala activity or activity patterns distributed across other brain systems for affective processes that are likewise implicated in CVD risk, such as the cingulate cortex, medial prefrontal cortex, insula, and cell groups of the brainstem (Gianaros & Jennings, 2018; Kraynak, Marsland, & Gianaros, 2018; Oppenheimer & Cechetto, 2016).

In the latter regard, a limitation to most prior studies of the neural correlates of CVD risk is that they have included small samples that constrain attempts to cross-validate candidate brain phenotypes that are reliably expressed by individuals. This issue is especially important insofar as many conventional and ostensibly affective neuroimaging tasks may not reliably index

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trait-like individual differences in stimulus-evoked brain activity (Elliott et al., 2020; Infantolino, Luking, Sauder, Curtin, & Hajcak, 2018; Nord, Gray, Charpentier, Robinson, & Roiser, 2017). As a result, brain metrics derived from such tasks in the context of small sample sizes may exhibit variable or unreliable relationships with other individual difference measures, such as CVD risk markers. Lastly, neuroimaging studies of CVD risk often use mass-univariate or potentially biased region-of-interest (ROI) quantitative methods that further engender issues with (a) multiple statistical testing, (b) reproducibility, and (c) accurately identifying whole-brain, distributed (multivariate), and generalizable neural patterns that predict cardiovascular risk variables—beyond localized patterns from just a single brain region (e.g., the amygdala) (Woo, Chang, Lindquist, & Wager, 2017).

Accordingly, the present study applied whole-brain, multivariate, machine-learning, and cross-validation methods to test whether individual differences in a vascular indicator of preclinical atherosclerosis and CVD risk (carotid artery intima-media thickness; CA-IMT) are reliably predicted by distributed brain activity patterns assessed by functional magnetic resonance imaging (fMRI) during three affective information processing tasks. As a preclinical marker of CVD risk, CA-IMT increases from a progressive thickening of the intimal and medial layers of the carotid arterial wall (Stein et al., 2008), and it has been widely used to measure the atherosclerotic disease process (Oygarden, 2017; Wendell, Waldstein, Evans, & Zonderman, 2017). CA-IMT associates with pathological and histologic measures of atherosclerosis in autopsy studies (Pignoli, Tremoli, Poli, Oreste, & Paoletti, 1986; Wong, Edelstein, Wollman, & Bond, 1993), and it has been shown to predict future CVD events (Baber et al., 2015; Peters, den Ruijter, Bots, & Moons, 2012; Polak, Szklo, & O'Leary, 2017). Two of the three affective fMR) tasks were standard matching-to-sample tasks, in which participants matched human faces depicting various emotions. The final task was an emotion processing and responding task, in which participants viewed neutral and unpleasant images from the International Affective Picture System (IAPS; Lang, Bradley, & Cuthbert, 2008) and provided subjective

ratings of experienced negative affect. Using three affective fMRI tasks enabled tests of whether any observed predictive associations between brain activity and CA-IMT would generalize across a broader range of affective cues than used previously (Gianaros et al., 2008). And in contrast to prior studies that relied on mass-univariate or ROI-based approaches to identify affective neural correlates of CVD risk, the present study used multivariate and whole-brain analysis methods that combined predictor dimensionality reduction, penalized linear regression, and cross-validation as per recent approaches and recommendations for predictive and patternbased modeling (e.g., Eisenbarth, Chang, & Wager, 2016; Gianaros, Sheu, et al., 2017; Poldrack, Huckins, & Varoquaux, 2019; Scheinost et al., 2019; Woo et al., 2017). Specifically, patterns of brain activity that were evoked by the affective tasks served as multivariate predictor variables, with CA-IMT serving as the outcome variable. Using nested cross-validation, the multivariate neural predictors of CA-IMT were tested for generalizability among study participants whose data were not used for predictive model building and pattern identification. Secondary analyses tested the (a) specific contribution of the amygdala in explaining variance in CA-IMT—owing to its prominence in past literature, (b) influences of age, sex, and other known cardiovascular risk factors, and (c) psychometric properties of brain activity changes evoked by the affective tasks

Materials and Methods

Participants

Data reported herein were derived from baseline (cross-sectional) assessments of participants from two studies: the Adult Health and Behavior project – Phase 2 (AHAB-2) and the Pittsburgh Imaging Project (PIP). Both AHAB-2 (N = 490) and PIP (N = 331) are registries of behavioral, biological, and neural correlates of CVD risk among otherwise healthy community-dwelling adults (aged 30-54 years). Across both cohorts, there was an approximate balance of men and women (AHAB-2: 53% women; PIP: 50% women in the full study samples). All participants provided informed consent. The University of Pittsburgh Human Research

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Protection Office granted approval for AHAB-2 (Protocol ID: 07040037) and PIP (07110287), as well as their aggregation to create a common data registry (19030174). The online Supplementary material provides study details.

Overview of study protocols

AHAB-2 and PIP entailed similar medical interviews, as well as multiple visits to assess sociodemographic factors, preclinical atherosclerosis (see below), and CVD risk. The latter included seated resting BP measured over multiple visits, waist circumference, and fasting glucose, high-density lipoprotein cholesterol (HDL), total cholesterol, and triglycerides, as detailed for AHAB-2 (Gianaros et al., 2014) and PIP (Gianaros, Sheu, et al., 2017).

AHAB-2 participants completed two versions of an fMRI facial-expression, matching-tosample task (see below). Sub-groups of AHAB-2 and PIP participants also completed the same affective processing and responding task, which involved viewing IAPS stimuli (see below). Group and individual-level fMRI data from the present report are available at NeuroVault (see Supplementary data; Gorgolewski et al., 2015).

Preclinical atherosclerosis

AHAB-2 and PIP participants underwent CA-IMT assessments by trained sonographers using an Acuson Antares ultrasound device (Acuson-Siemens, Malvern, PA). Details on CA-IMT assessments are online (see Supplementary material). All CA-IMT values obtained for the left and right carotid arteries were averaged bilaterally to compute a grand mean CA-IMT, which served as the primary outcome variable.

The test-retest reliability of CA-IMT was determined in 14 participants in PIP who underwent ultrasonography twice (median 90 days apart, range 42–153 days). Across occasions, CA-IMT differed by -0.09 to 0.06mm (M = -0.005, SD = 0.042mm). The intraclass coefficient (ICC; ratio of variability between participants / total variability) was 0.82, indicating good reliability.

Facial-expression matching tasks

AHAB-2 participants completed two versions of a facial-expression matching task. The first was identical to the task used in our prior report on amygdala reactivity and CA-IMT (Gianaros et al., 2008). In this task, participants completed 4 blocks of a facial-expression matching-to-sample condition, which was interleaved with 5 blocks of a shape-matching (sensorimotor) control condition. In the facial-expression matching condition, participants saw 3 same-sex faces in an array for each trial, all expressing either fear or anger. Participants chose 1 of 2 faces at bottom that was identical to a center target at top. Each block consisted of 6 trials (3 fear, 3 anger; 3 all-male, 3 all-female), and each trial lasted for 4 sec (variable 1.5, 3.5, and 5.5 sec variable inter-trial interval; ITI). All images were in black and white and were drawn from the Pictures of Facial Affect (PFA) stimulus set (Ekman & Friesen, 1976). In the control condition, participants also matched-to-sample, but instead used images of circles, vertical ellipses, and horizontal ellipses. Each trio of shapes was shown for 4 sec, with a 2 sec ITI. The total task length was 6 min and 36 sec, including 6 sec for initial magnetic equilibration.

In the second version of the task, participants completed 4 blocks of facial-expression, matching-to-sample trials (2 sec trials; 60 sec block length; variable ITI = 1, 3 and 5 sec) and 5 interleaved blocks of sensorimotor control trials (2 sec trials; 48 sec block length; ITI = 2 sec). In this version, however, blocks of trials were comprised of trios of angry, fearful, happy, or neutral facial expressions (12 trials per block). All 3 faces for a given trial were matched for sex and emotional expression. Images in this task were shown in color and drawn from the NIM-STIM Set of Facial Expressions (2006). The total task length was 9 min and 12 sec, including 6 sec for initial magnetic equilibration. In both tasks, each block began with a 2 sec cue: "Match Faces" or "Match Shapes."

IAPS task

Sub-groups of AHAB-2 and PIP participants completed an affective processing and responding task, which has been detailed previously (Gianaros et al., 2014). Participants saw

30 unpleasant and 15 neutral IAPS images (Lang et al., 2008). Participants were first trained and then instructed to (a) "Look" and attend to images or (b) "Decrease" and change their thinking about the image to feel less negative. Trials were comprised of a 2-sec cue ("Look" or "Decrease"), followed by a 7-sec IAPS image presentation. After image viewing, participants rated their emotional state ("How negative do you feel?") on a 5-point Likert-type scale in a 4-sec rating period (1 = neutral, 5 = strongly negative). A variable (1-3 sec) rest period followed each rating period. The entire task duration was 11 min and 16 sec (15 'Look neutral' trials; 15 'Look negative' trials; 15 'Decrease negative' trials). Images were presented such that no more than 2 identical trial types ("Look negative" or "Decrease negative") were consecutive, and no more than 4 unpleasant images were consecutive. Given the focus of this report (and for comparability to the facial-expression tasks involving the passive viewing of affective stimuli), only 'Look neutral' and 'Look negative' trials overlapped by 87% (13/15 shared images) and neutral images by 13% (2/15 shared images).

As a manipulation check, we verified that unpleasant IAPS stimuli increased selfreported negative affect (see Table 1), t(337) = 61.80, p < 0.001. The difference in negative affect ratings between 'Look negative' and 'Look neutral' trials (i.e., self-reported affective reactivity) was not statistically correlated with CA-IMT, r(336) = -0.02, p = 0.76. These behavioral data were thus omitted from further CA-IMT analyses.

E-Prime software (Psychology Software Tools, Sharpsburg, PA) was used to present stimuli and record behavioral responses in the facial-expression and IAPS tasks. Table S1 lists image stimuli used for all tasks (see online Supplementary material).

MRI data acquisition and analysis

Functional blood oxygenation level-dependent (BOLD) images from AHAB-2 and PIP participants were collected on the same 3 Tesla Trio TIM whole-body scanner (Siemens,

Erlangen, Germany), equipped with a 12-channel head coil. Data acquisition and preprocessing details are online (see Supplementary material).

In within-individual fMRI analyses, univariate general linear models (GLM's) were estimated to compute condition or event contrast maps that were later used for prediction analysis. Specifically, task conditions (or events) were modeled by rectangular waveforms convolved with the default hemodynamic response function in SPM12. For the facial-expression matching tasks, these regressors modeled blocks of face- and shape-matching conditions. For the IAPS task, these regressors modeled events of the trial (i.e., cue, picture, rating period, rest). In each GLM, the 6 realignment parameters from preprocessing were included as nuisance regressors, and low-frequency artifacts were removed by a high-pass filter (128 sec). Error variance was estimated and then weighted by restricted maximum likelihood estimation, as implemented in the RobustWLS toolbox, v4.0 (Diedrichsen & Shadmehr, 2005). Linear contrasts were then computed, including the '*Faces vs. Shapes*' contrast for the facial-expression matching tasks and '*Look negative* vs. *Look neutral*' contrast for the IAPS task. Main effects of the tasks were determined from these GLM's at the group (between-individual) level with whole-brain correction for multiple testing using a voxel-wise and false discovery rate threshold (0.05) combined with a voxel-extent threshold of k = 50.

Missing data

Of the 490 AHAB-2 participants, 430 had complete fMRI data from the facial-expression task using PFA faces and 453 using NIM-STIM faces. Of the 430 with complete PFA data, 1 was excluded because of an anatomical abnormality, and 2 were excluded because of missing CA-IMT. Of the 453 with complete NIM-STIM data, 2 were excluded because of missing CA-IMT, and 1 was excluded for insufficient brain coverage. AHAB-2 analyses were thus conducted on 427 participants with PFA data and 450 participants with NIM-STIM data.

There were 180 AHAB-2 and 174 PIP participants with IAPS fMRI data. As noted, 15 people participated in AHAB-2 and PIP, and we used only their AHAB-2 data. Of the 354 total

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In interim, there were 821 (N = 490 AHAB-2, N = 331 PIP) total individuals, but only 617 (N = 458 AHAB-2, N = 159 PIP) had complete and useable data for analysis, including useable fMRI data for one or more of the 3 affective tasks reported here. Table 1 summarizes demographic and other characteristics of the AHAB-2 and PIP analytic samples.

Multivariate prediction of CA-IMT

We used principal component regression combined with least absolute shrinkage and selection operator (hereafter, LASSO-PCR), as introduced for fMRI by Wager and colleagues (2011; 2013). LASSO-PCR was used to specifically test whether CA-IMT was predicted by BOLD signal responses evoked by the facial-expression matching tasks ('Faces vs. Shapes') and by the IAPS task ('Look negative vs. Look neutral'). To this end, evoked responses within gray matter voxels from all participant contrast maps served as the multivariate predictor, with CA-IMT as the dependent (outcome) variable. LASSO-PCR consisted of the following steps. To reduce voxel-wise predictor dimensionality and account for multicollinearity across voxels, the multivariate voxel-wise predictor matrix was first submitted to singular value decomposition (SVD). CA-IMT was then regressed onto the remaining principal components by LASSO using the 'lassoglm' function in MATLAB (Hastie, Tibshirani, Friedman, & Franklin, 2005; Tibshirani, 1996). In LASSO, estimated beta coefficients are subject to the L1 penalty (regularization). Applying this penalty, whose effect is modulated by the parameter lambda (λ), shrinks (penalizes) non-significant beta coefficients to zero to select predictive features (Efron, Hastie, Johston, & Tibshirani, 2004; Zhao, Rocha, & Yu, 2009). Following feature selection, nonzero coefficients ($\hat{\beta}$) are projected back to voxel space via the SVD transformation matrix to produce a whole-brain weight map, \widehat{w} . The dot product of this whole-brain weight map (\widehat{w}) and a participant's contrast map of evoked signal change generates a predicted estimate of CA-IMT.

Nested and k-fold cross-validation was implemented to first optimize the shrinkage parameter λ in an 'inner loop' and then determine predictive generalizability of the multivariate weight maps in an 'outer loop'. Hence, to first identify the optimal shrinkage parameter, λ , and thus the number of principal components that best predicted CA-IMT, we used a 5-fold cross-validation in the inner loop. In this way, participants were divided into training and validation samples by stratifying over the CA-IMT distribution. Within each cross-validation fold, LASSO-PCR was conducted on training samples using a sequence of 1000 λ values, and the performance of each λ was evaluated by calculating the average of the mean squared error (MSE) between predicted and observed CA-IMT across the validation samples. After identifying the optimal λ , the entire LASSO-PCR procedure was repeated using this λ on the entire inner loop sample, producing a whole-brain multivariate predictive pattern of CA-IMT.

To test the generalizability of the LASSO-PCR models and to generate unbiased estimates of predicted CA-IMT for each participant, we repeated the above process in an outer cross-validation loop. For the two facial-expression tasks in the AHAB-2 cohort, the outer loop comprised a 5-fold cross-validation. For the IAPS task, the outer loop comprised a 2-fold cross-validation, which separated training and validation sets according to study membership (i.e., AHAB-2 vs. PIP). The latter strategy was used to conservatively test for generalizability from one study sample to the other in view of observed cohort differences in CA-IMT (M CA-IMT in AHAB-2 = 0.638 mm; M CA-IMT in PIP = 0.609 mm; t(336) = 2.10, p = 0.04), prevalence of non-overlapping IAPS images (see Methods), and fMRI spatial signal coverage (mean percent coverage of all gray matter voxels in AHAB-2 = 86.27%; mean percent coverage in PIP = 93.59%, t(336) = 29.83, p < 0.001). Thus, the latter cross-validation procedure attempts to "replicate" IAPS-based prediction findings for CA-IMT across the two cohorts (between AHAB-2 and PIP). Complete results and illustrations are available online (see Supplementary material).

The final predictive performance of a given weight map was summarized using several metrics describing the association between predicted and observed CA-IMT values

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concatenated across the k folds of the outer loops (Forman & Scholz, 2010). Here, the similarity between predicted and observed values was summarized by the Pearson correlation coefficient and its corresponding 95% confidence interval (CI) and p-value. The discrepancy between predicted and observed values was calculated using mean absolute error (MAE). Following guidelines for predictive modeling (Poldrack et al., 2019; Scheinost et al., 2019), variance in observed values explained by predicted values (R²) was calculated by inserting the out-of-sample predictions in the sums-of-squares formulation. To aid in interpreting results that did not meet conventional statistical significance, we computed Bayes factors (BF10, BF01) using the 'BayesFactors' package (v0.9.12-4.2; Morey & Rouder, 2011) in R (R-Core-Team, 2014). BF10 values reflect the probability of the alternative hypothesis, relative to the null. BF01 values reflect the inverse of BF10 values, corresponding to evidence for the null relative to the alternative hypotheses.

To identify which particular voxels reliably contributed to a given multivariate weight map, we used bootstrap resampling (1000 iterations) and repeated the LASSO-PCR and k-fold cross-validation steps. P-values for voxel-wise estimates were then computed by nonparametric percentile methods, and the bootstrapped multivariate weight maps were thresholded at p < 0.05 (2-tailed) with an extent threshold of k = 50 voxels.

LASSO-PCR and nested cross-validation procedures were implemented using in-house MATLAB code adapted from Wager and colleagues (Wager et al., 2011; Wager et al., 2013). All code is available online (see Supplementary material).

Ancillary tests of amygdalar contributions

Amygdala functionality has been implicated in CVD risk (see Introduction). This raises the question of whether amygdala activity is necessary or sufficient to predict CA-IMT. To address this question, we repeated the above nested cross-validated LASSO-PCR steps for each affective task, first restricting all predictor voxels to a bilateral anatomical mask of the amygdala. Next, we performed a 'virtual lesion' analysis, wherein all voxels of the amygdala mask were excluded from nested cross-validation. The anatomical mask included the left and right amygdalae of the Automated Anatomical Labeling atlas (Tzourio-Mazoyer et al., 2002) in the Wake Forest University PickAtlas Toolbox (Maldjian, Laurienti, Kraft, & Burdette, 2003).

Ancillary tests of discriminant prediction

We further explored whether CA-IMT could be predicted by an entirely different brain pattern that statistically explains self-reports of negative affect evoked by unpleasant IAPS stimuli; namely, the Picture-Induced Negative Emotion Signature (PINES; Chang, Gianaros, Manuck, Krishnan, & Wager, 2015). If so, then this could suggest that brain patterns derived here are not uniquely predictive of preclinical atherosclerosis, but rather interchangeable or redundant with brain correlates of self-reported negative affect.

Ancillary tests of 'confounders'

Demographic, anthropometric, and biological factors were explored as effect moderators or 'confounders' that may impact primary study findings. These included tests of age and sex as moderators, as well tests of age, sex, and components of the metabolic syndrome as covariates. For the latter and following prior work (Gianaros, Kuan, et al., 2017), a standard composite cardio-metabolic risk (CMR) score was computed by z-scoring systolic BP, waist circumference, and fasting concentrations of glucose, HDL levels, and triglycerides and then averaging the z-scores (HDL signs reversed).

Psychometric properties of task-related fMRI activity

To assess the internal consistency of fMRI activity changes evoked by tasks, we executed the split-half method in line with procedures of Infantolino and colleagues (2018). Specifically, for each participant, GLM analyses were estimated and contrast images were generated separately using approximate partitions of the first and second half of the BOLD signal data from each task. For the PFA-based faces task, there were four expression-matching blocks and five shape-matching blocks. The last block of shape-matching trials was not included, so that the first and second partitions of the data each had two face- and shape-

matching blocks. For the NIM-STIM-based faces task, there were four expression-matching blocks and five shape-matching blocks. Again, the last block of shape-matching trials was not modelled. For the IAPS task, there were 45 trials that could not be partitioned equally in a conventional split-half analysis because of the nature of the event-related design and its stimulus ordering. As a result, the fMRI run was partitioned into the first 23 trials and the last 22 trials (resulting in 10 vs. 5 'Look negative' and 7 vs. 8 'Look neutral' trials, respectively).

We first computed the split-half, internal consistency of amygdala activity within each task by extracting the average contrast value within the bilateral AAL amygdala mask, and we examined the reliability of the extracted parameter estimates using the Spearman-Brown (SB) correction method. For completeness of reporting, we then generated voxel-wise internal consistency maps for each task. These voxel-wise maps are available online at NeuroVault (see also Supplementary material). Finally, we computed the internal consistency of the weight maps generated by LASSO-PCR to predict CA-IMT. To this end, we computed the dot-product between each whole-brain weight map (\widehat{w}) with all pairs of participant contrast maps generated from the split-half procedure, which effectively tested for the consistency of CA-IMT prediction within participants (Woo & Wager, 2016).

Results

Main effects of the facial-expression matching tasks

The 'Faces vs. Shapes' contrasts for the facial-expression matching tasks using stimuli from the PFA and NIM-STIM sets showed that both tasks engaged an ensemble of cortical and subcortical brain regions in the AHAB-2 cohort (see Figure 1a and 1b and Supplemental Tables S2-S3). Although both tasks positively engaged the amygdala on average (Figure 1a, 1b), the internal consistency metrics for extracted amygdala activation values were poor for tasks using PFA (0.27) and NIM-STIM (0.11) facial expressions. The latter observations match those of prior reports indicating poor psychometric properties of extracted amygdala activation values, as evoked by facial-expression matching-to-sample tasks (Elliott et al., 2020; Infantolino et al.,

2018). However, we note that on a voxel-wise basis, internal consistency values varied appreciably across the brain for both face-matching tasks, with some areas (e.g., in visual cortex) exhibiting moderate-to-good internal consistency (interactive maps available at NeuroVault; see also Supplementary material).

Main effects of the IAPS task

The 'Look negative vs. Look neutral' contrast revealed that viewing unpleasant IAPS images also engaged a distributed ensemble of brain regions known to respond to complex visual affective scenes, including a positive engagement of the amygdala and areas of the medial prefrontal, anterior cingulate, and anterior insular cortices. Figure 1c and Supplemental Table S4 summarize these main effects, as aggregated across the AHAB-2 and PIP cohorts (N = 338). Across participants, the internal consistency of extracted activation values from the amygdala were comparable to those for the facial-expression matching tasks (0.24). As observed for the facial-expression matching tasks, there was wide variability in internal consistency on a voxel-wise basis (maps available at NeuroVault; see also Supplementary material).

Prediction of CA-IMT from whole-brain patterns evoked by the facial-expression matching tasks

<u>PFA-Based Prediction.</u> Using cross-validated LASSO-PCR, we found that CA-IMT was predicted by a model trained on the contrast of 'Faces vs. Shapes' (inclusive of angry and fearful faces) from the PFA facial-expression matching task, optimal $\lambda = 0.0113$, 24 principal components retained, MSE = 0.0137. As shown in Figure 2a, nested cross-validation showed that CA-IMT predicted by the whole-brain multivariate brain pattern correlated with observed CA-IMT, r(425) = 0.13, 95% CI = 0.04 to 0.23, p = 0.005, MAE = 0.089, R² = 0.014, BF10 = 5.16, BF01 = 0.19. The internal consistency of the PFA-based weight map was 0.73. Figure 3a illustrates brain areas within the whole-brain multivariate pattern that consistently predicted CA-IMT across individuals in hold-out validation sample testing, as determined by bootstrapping.

We reiterate that a whole-brain pattern was used for cross-validated prediction; however, it is noteworthy from the lower right panel of Figure 3a that the predictive contributions of the amygdala were mostly *negative* (94 or 20% negative vs. 17 or 3.6% positive voxel weights out of 468 possible voxels within the amygdala mask). The latter indicates that there were more voxels within the amygdala where *lower* evoked activity predicted *higher* CA-IMT.

When cross-validated LASSO-PCR was restricted to voxels within the amygdala, CA-IMT continued to be predicted across individuals—albeit with a smaller effect size, r(425) = 0.10, 95% CI 0.003 to 0.19, p = 0.044, MAE = 0.010, R² = 0.006, BF10 = 0.84, BF01 = 1.19. And, in a 'virtual lesion' analysis in which we applied the amygdala mask to omit this region from the whole-brain LASSO-PCR and nested cross-validation steps, we continued to observe an association between predicted and observed CA-IMT, r(425) = 0.13, 95% CI 0.04 to 0.22, p = 0.007, MAE = 0.090, R² = 0.014, BF10 = 4.16, BF01 = 0.24. The latter findings suggest that the amygdala was not necessary for predicting CA-IMT from the whole-brain predictive pattern.

The relationship between predicted vs. observed CA-IMT was not moderated by age ($\beta = 0.04$, SE = 0.04, p = 0.340, BF10 = 0.16, BF01 = 6.25) or sex ($\beta = 0.06$, SE = 0.05, p = 0.187, BF10 = 0.43, BF01 = 2.33) when the latter were added as interaction terms. In tests of confounding influences, however, covariate adjustment for age, sex, and the composite cardiometabolic risk score demonstrated that predicted and observed CA-IMT no longer correlated at a conventional statistical threshold of p < 0.05, ($\beta = 0.03$, SE = 0.04, t(422) = 0.80, p = 0.424, BF10 = 0.14, BF01 = 7.15).

Lastly, in a test of discriminant prediction, applying the PINES brain pattern for self-reported negative affect to the PFA contrast maps of 'Faces vs. Shapes' failed to statistically predict CA-IMT, r(425) = -0.06, 95% CI -0.15 to 0.04, p = 0.22, MAE = 1.608, $R^2 < 0$, BF10 = 0.24, BF01 = 4.16. This suggests that multivariate brain patterns for predicting self-reported negative affect vs. CA-IMT are dissociable and not redundant for the PFA task paradigm.

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<u>NIM-STIM-Based Prediction.</u> Using cross-validated LASSO-PCR, we found that CA-IMT was not reliably predicted by whole-brain patterns revealed by the 'Faces vs. Shapes' contrast (inclusive of angry, fearful, happy, and neutral facial expressions) from the facial-expression matching task using NIM-STIM images. This null result was reflected in predicted CA-IMT values that did not statistically correlate with observed CA-IMT, r(448) = 0.04, 95% CI -0.06 to 0.13, p = 0.354, MAE = 0.088, R² < 0, BF10 = 0.17, BF01 = 5.88. Similar findings were revealed by post-hoc tests of whether the specific contrast of 'Fear and Anger vs. Shapes' predicted CA-IMT, r(448) = 0.07, 95% CI -0.02 to 0.16, p = 0.127, MAE = 0.088, R² < 0, BF10 = 0.34, BF01 = 2.88 (full results available online in the Supplementary material). These post-hoc tests were executed for a more direct comparison of findings from the PFA-based facial-expression paradigm, which included only expressions of fear and anger. Because a reliable whole-brain pattern of brain activity predictive of CA-IMT could not be generated using the NIM-STIM based task, we did not explore ancillary moderation, confounder, internal consistency, or PINES tests. **Prediction of CA-IMT from whole-brain activity patterns evoked by IAPS images**

Using cross-validated LASSO-PCR, we found that CA-IMT was predicted by a wholebrain multivariate pattern derived from the 'Look negative vs. Look neutral' contrast based on the IAPS task, optimal $\lambda = 0.0049$, 113 principal components retained, MSE = 0.0099. As revealed by nested cross-validation and shown in Figure 2b, predicted and observed CA-IMT values were statistically correlated, r(336) = 0.19, 95% CI = 0.10 to 0.28, p < 0.001, MAE = 0.076, R² = 0.023, BF10 = 47.10, BF01 = 0.02. The internal consistency of the multivariate IAPS pattern was 0.82. Figure 3b shows brain areas within the multivariate pattern that consistently predicted CA-IMT, as determined by bootstrapping. We note that surviving predictive weights within the amygdala were uniformly negative (90 or 19% negative vs. 0 positive voxel weights out of 468 possible voxels), again indicating that *lower* evoked activity predicted *higher* CA-IMT.

When the cross-validated LASSO-PCR procedure was restricted to the anatomical mask of the amygdala, the LASSO-PCR model trained on one study cohort failed to generalize to predict CA-IMT in the other study cohort. Moreover, after excluding voxels within the amygdala from LASSO-PCR in a 'virtual lesion' analysis, we continued to observe an association between predicted and observed CA-IMT across participants, r(336) = 0.19, 95% CI = 0.10 to 0.28, p = 0.001, MAE = 0.076, R² = 0.022, BF10 = 45.70, BF01 = 0.02. The latter findings indicate that amygdala activity was neither necessary nor sufficient to predict CA-IMT.

In a test of discriminant prediction, applying the PINES pattern for predicting selfreported negative affect to 'Look negative vs. Look neutral' contrast maps failed to predict CA-IMT, r(336) = 0.04, 95% CI -0.07 to 0.15, p = 0.442, MAE = 1.023, $R^2 < 0$, BF10 = 0.17, BF01 = 5.88, again suggesting that brain patterns for self-reported negative affect vs. CA-IMT are dissociable.

The relationship between predicted and observed CA-IMT was not moderated by age (β = 0.07, SE = 0.05, p = .155, BF10 = BF10 = 0.33, BF01 = 3.03) or sex (β = -0.05, SE = 0.05, p = .384, BF10 = 0.29, BF01 = 3.45). And in contrast to findings from the facial-expression task using PFA stimuli, adjustment for age, sex, and the composite index of cardiometabolic risk demonstrated that predicted and observed CA-IMT continued to correlate, β = 0.17, SE = 0.04, t(333) = 3.84, p < 0.001, BF10 = 147.81, BF01 < 0.01.

Discussion

This study tested whether whole-brain fMRI activity patterns evoked by three affective information processing tasks would reliably predict individual differences in CA-IMT, a vascular marker of preclinical atherosclerosis and CVD risk. First, as revealed by multivariate, machine-learning and cross-validation methods, ~1.4% of the variance in CA-IMT across individuals was explained by a whole-brain pattern that was trained on fMRI activity elicited by viewing facial expressions of emotion (anger and fear). These expressions were presented in a canonical matching-to-sample task using PFA stimuli (Figures 2a, 3a). Although the latter predictive relationship between fMRI activity and CA-IMT was observed in a large cohort of individuals (N

= 427) through k-fold cross-validation, it failed to persist after accounting for age, sex, and known CVD risk factors. Second, CA-IMT was not reliably predicted by a whole-brain pattern elicited by viewing angry, fearful, happy, and neutral facial expressions in a similar matching-tosample task using NIM-STIM stimuli (N = 450; see online Supplement). Comparable findings were observed when models were trained only on fMRI activity patterns evoked by angry and fearful NIM-STIM facial expressions (see online Supplement). Third, ~2.3% of the variance in CA-IMT was explained by a whole-brain pattern trained on fMRI activity elicited by viewing IAPS images of unpleasant affective scenes (see Figures 2b, 3b). The latter predictive relationship (a) generalized across two different study samples (N's of 179 in AHAB-2 and 159 in PIP) and (b) persisted after accounting for age, sex, and known risk factors for CVD. Fourth, aggregate evidence across all predictive models was largely inconsistent with (or opposite to) prior findings suggesting that greater amygdala reactivity to affective information relates to greater preclinical atherosclerosis, as measured by CA-IMT (Gianaros et al., 2008). Rather, multivariate associations between amygdala activity and CA-IMT were largely negative when present. Lastly, CA-IMT was not statistically associated with self-reports of negative affect induced by viewing unpleasant affective scenes or a previously-established brain pattern (i.e., the PINES) that predicts self-reported negative affect (Chang et al., 2015), suggesting dissociable neural correlates of self-reported negative affect and vascular markers of CVD risk. Collectively, the present findings provide novel evidence for neural correlates of CVD risk that extend beyond localized amygdala activity to include multivariate and generalizable patterns that are distributed across other brain systems for affective processing. As described below, the interpretation and predictive strength of these multivariate neural correlates may depend on their psychometric properties, the influence of other cardiovascular risk factors, and perhaps the specific affective cues used to evoke neural activity.

To elaborate, recent evidence suggests that amygdala functionality may associate with clinically-relevant cardiovascular endpoints and surrogate markers of CVD risk (e.g., Ishai et al.,

2019; Osborne et al., 2019; Tawakol et al., 2017). Predating the latter evidence were findings that greater amygdala activity to facial expressions of fear and anger was associated with a greater severity of preclinical atherosclerosis, suggesting that amygdala 'hyperactivity' evoked by affective cues (facial expressions of emotion) may comprise a neural correlate of CVD risk (Gianaros et al., 2008). Here, we administered the same facial-expression matching-to-sample fMRI task that we used previously, wherein participants matched faces depicting anger and fear derived from the PFA stimulus set. This task was complemented by another facial-expression matching task that included not only angry and fearful facial expressions, but also happy and neutral expressions that were all drawn from a different and more recently developed stimulus set (NIM-STIM). The latter NIM-STIM-based task was administered in an attempt to test whether any findings derived from the PFA-based task generalized to a broader range of affectively salient expressions and human models.

Using the PFA-based task, we found that a *whole-brain* activity pattern predicted CA-IMT in final cross-validation testing (R² = 0.014). This cross-validated and predictive relationship, however, did not depend on the contributions of the amygdala in a 'virtual lesion' analysis. Nor did the predictive relationship derived from the whole-brain pattern remain statistically significant following covariate control for age, sex, and known cardiovascular risk factors. In addition, the predictive weights within the amygdala in both whole-brain and ROI-based analyses were largely *negative*, not positive—meaning that *greater* CA-IMT was predicted by *lower* levels of evoked amygdala activity. The latter observations disagree with the positive associations we previously reported between amygdala activation to angry and fearful facial expressions and CA-IMT using mass univariate and ROI methods (Gianaros et al., 2008). Moreover, the NIM-STIM-based task failed to produce a generalizable multivariate brain pattern that predicted CA-IMT using either all facial expressions or only a subset expressing anger and fear. One reason why fMRI activity from these two facial-expression matching tasks did not consistently relate to CA-IMT may be due to their relatively limited psychometric properties and,

thus, unclear suitability for individual difference research. The latter conclusion is supported not only by the present findings, but also those reported by Infantolino and colleagues (2018) and summarized by meta-analysis (Elliott et al., 2020). Indeed, post-hoc exploratory findings (Figure S5) suggested that internal consistency values were related to predictive performance on a voxel-wise basis across all tasks used in the present study. As specifically shown in Figures S5 (panels e and f), voxels exhibiting lower internal consistency values tended to exhibit lower prediction weights. Consequently, internal consistency appeared to "constrain" prediction weights in a manner that accords with classical test theory (Cohen, Cohen, West, & Aiken, 2003). In these regards, the psychometric characteristics of fMRI activity patterns evoked by affective and other task paradigms should be considered in individual differences research on CVD risk.

Compared to findings from the facial-expression matching tasks, a LASSO-PCR model trained on whole-brain activity evoked by viewing complex and unpleasant IAPS stimuli in one study cohort generalized to predict CA-IMT in a different study cohort, with a relatively larger coefficient-of-determination (R² = 0.023). Viewing unpleasant IAPS stimuli also engaged a broad ensemble of brain areas implicated in affective processing and visceral control functions relevant to CVD risk, including the insula, hypothalamus, brainstem, and areas of the anterior cingulate and medial prefrontal cortices (see Figure 1c). Moreover, the multivariate (i.e., whole-brain) predictive pattern of activity evoked by the IAPS task exhibited reasonable psychometric properties, particularly relative to those observed for the facial-expression matching tasks. Indeed, the multivariate IAPS pattern predictive of CA-IMT exhibited an internal consistency value of 0.82, approximating the test-retest reliability (ICC) value observed in one of our samples (PIP) for CA-IMT itself (0.82). It is notable that this predictive model continued to account for CA-IMT above-and-beyond canonical CVD risk factors in final cross-validation testing (age, sex, and a composite cardio-metabolic risk score).

Also notable from Figure 3b is that regions most consistently predictive of CA-IMT within the whole-brain pattern encompassed those implicated in linking affective processes (and negative affect in particular) to CVD risk via downstream influences over health behaviors and systemic physiology (Kraynak, Marsland, & Gianaros, 2018), including predictive contributions from the insula, hypothalamus, brainstem, and areas of the anterior cingulate and medial prefrontal cortices (Figure 3b; Supplementary Table S6). However, the amygdala appeared neither necessary nor sufficient for predicting CA-IMT across individuals in ROI and "virtual lesion" analyses within the context of the IAPS task. Lastly, viewing unpleasant IAPS stimuli induced a mean increase in self-reported negative affect across individuals. These self-reports, however, were insufficient as explanatory predictors of CA-IMT, as was a multivariate brain pattern that predicts self-reported negative affect (i.e., the PINES pattern). These observations again suggest a dissociation between the neural correlates of the subjective experience of negative emotion and the neural correlates of CVD risk (e.g., as measured at the level of the peripheral vasculature by CA-IMT). Viewing complex and unpleasant affective scenes may thus not only induce negative affect, but also a unique pattern of brain activity that exhibits a predictive relationship (albeit of small effect size) with CA-IMT.

Inferences from the present findings are constrained by study design and sample characteristics, as well as related methodological limitations. First, our cross-sectional observations preclude causal inference. For example, in a health neuroscience framework (Erickson, Creswell, Verstynen, & Gianaros, 2014), it may be that the IAPS-based multivariate brain pattern that predicted CA-IMT corresponds to a brain phenotype that reflects a link between negative affective processes and CVD risk via top-down (brain-to-body or brain-to-behavior) pathways. These pathways could encompass individual differences in affect-related neural influences on peripheral physiology and health behaviors. At the same time, preclinical cardiovascular pathophysiology could plausibly influence patterns of brain activity that predicted CA-IMT via bottom-up (body- or behavior-to-brain) pathways, such as via the ascending

(afferent) influences of immune and autonomic input to the brain that in turn bias how individuals process and respond to affective cues (Kraynak, Marsland, & Gianaros, 2018; Kraynak, Marsland, Wager, et al., 2018). Second, the study samples were predominately of European ancestry, well-educated, and limited to people free of major chronic illnesses and medication regimens that could have confounded interpretations of CA-IMT and related CVD risk metrics. As a result, the relevance of our findings to demographically diverse individuals and clinical populations is unclear. However, our whole-brain maps are publicly available and could be applied in other populations and tested as predictors of clinical outcomes or other markers of preclinical CVD. Third, we adopted an omnibus training-and-testing methodology, collapsing across facial expressions of diverse emotions and IAPS stimuli depicting a range of affective scenes to produce contrast images of 'Faces vs. Shapes' and 'Negative vs. Neutral' activity patterns. This was done without an a priori focus on affect- or content-specific predictive testing. Using publicly available contrast images from the present study, it would be feasible for others to generate and test more granular affective predictive maps for CA-IMT. Fourth, it may be that metrics of resting metabolic activity within brain systems for affective processes or even distributed network (e.g., intrinsic, resting-state, or task-based connectivity) metrics may hold methodological and possibly psychometric advantages than the task-based activation metrics employed here, as indicated by work on CVD risk and resting metabolic activity within the amygdala (Tawakol et al., 2017). Finally, we may have observed differential predictions of CA-IMT across tasks because of design characteristics and specific affective cues. Both the face processing and IAPS tasks involved the passive encoding of salient affective cues; however, only the IAPS task employed complex emotional and naturalistic scenes that would plausibly evoke an affective experience (e.g., as verified in the present study by a mean increase in negative affect).

Other design differences may likewise explain divergent findings across the faceprocessing and IAPS tasks. For example, both face-processing tasks relied on blocked stimulus presentations, with repeated exposures to some cues (or human models) that were presented for 2 – 4 sec. By contrast, the IAPS task used an event-related design, with no repeated exposures to affective cues and longer stimulus presentation times (7 sec). In addition to differences in stimulus content, novelty, and the complexity of the affective cues (faces vs. scenes) there were notable differences in the control conditions used to create fMRI activation maps (shapes vs. scenes). Together, these paradigmatic differences could have impacted how affective cues were encoded and experienced, thus impacting manifest psychometric characteristics of relative changes fMRI activity patterns used for predictive modeling. In extension, it is notable that metrics reflecting individual differences in the habituation of taskevoked neural responses to repeated or sequential affective cues (e.g., within the amygdala) may hold promise for identifying replicable brain phenotypes for cardiovascular risk relative to conventional and directional activity changes studied here (Plichta et al., 2014). An important future direction will be to validate task paradigms and fMRI activity metrics that maximize psychometric properties and utility for individual difference research.

To close, individual differences in emotion and affect have been studied in association with risk for CVD, historically at the level of peripheral physiology, behavior, and self-report. This line of inquiry has been extended to the level of the brain in health neuroscience, especially within the past decade. While there has been a strong conceptual and empirical focus on the amygdala as key brain system for CVD risk, the present findings support a broader focus on distributed neural systems in association with markers of CVD risk and early (preclinical) pathophysiology. The present findings also underscore the relevance of multivariate and cross-validation methodologies, as well as psychometric characteristics of fMRI tasks, for building predictive models that characterize replicable affective neural correlates of CVD risk.

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Supplementary data

Supplementary data are available at SCAN online. Code for analysis, as well as group- and individual-level fMRI data are available at GitHub (https://github.com/CoAxLab/IMT_manuscript) and NeuroVault (https://neurovault.org/collections/EBPXRZTX/).

Conflict of Interest. None declared.

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Captions

<u>Figure 1.</u> Color-scaled T-maps of brain areas exhibiting significant BOLD signal changes for the contrasts of Faces vs. Shapes and Look negative vs. Look neutral are shown for the facial-expression matching task employing PFA (A) and NIM-STIM faces (B), as well as IAPS images (C), respectively. Maps in A - C correspond to statistical parametric T-maps, shown at a false discovery rate (FDR) threshold of 0.05 to control for multiple statistical testing across voxels, with an extent threshold of k > 50 voxels. Axial planes correspond to z coordinates -16, -8, and 0. Warmer colors (orange/yellow) reflect activation, whereas cooler colors (blue) reflect deactivation. Clusters are described in the online Supplementary Material.







Figure 2. A comparison of predicted vs. observed CA-IMT using cross-validated LASSO-PCR in the (A) PFA-based facial-expression matching task and (b) IAPS task.



<u>Figure 3.</u> Whole-brain weight maps used to predict CA-IMT from the (a) PFA-based facialexpression matching task and (b) IAPS task. The maps are thresholded to display voxels which reliably contributed to prediction. Here, the maps are shown at a threshold of p < .05 and k > 50*voxels*, based on bootstrapping (1000 iterations). Axial planes correspond to z coordinates -16, -8, and 0. Warmer colors (orange/yellow) reflect a positive pattern weight, whereas cooler colors (blue) reflect a negative weight.



	AHAB-2 <i>N=</i> 458	PIP N=159
Characteristic		
Women (%)	52.4	50.3
Age (years)	42.6 ± 7.4	39.7 ± 6.4
Race (%) Caucasian/White African American Multi-racial/ethnic	81.8 15.9 2.2	66.0 28.3 5.6
Number of school years completed	16.9 ± 2.9	16.2 ± 3.3
Waist circumference (in)	35.5 ± 5.5	35.2 ± 5.3 ⁶
Total cholesterol (mg/dL)	199.8 ± 38.3 ¹	182.8 \pm 33.0 7
Triglycerides (mg/dL)	109.4\± 68,2 ¹	$95.5\pm56.1\ ^{7}$
HDL (mg/dL)	55.9 ± 15.0 ¹	52.2 ± 16.0 7
Glucose (mg/dL)	98.3 ± 11.2 ¹	$\textbf{88.6}\pm\textbf{8.8}^{\text{7}}$
Clinic resting SBP (mmHg)	115.2 ± 11.1 ²	120.1 ± 10.5
Clinic resting DBP (mmHg)	72.2 \pm 8.1 2	$\textbf{72.0} \pm \textbf{8.6}$
Accuracy in PFA faces task (%)	$96.0\pm14.0~^3$	-
Accuracy in NIM-STIM faces task (%)	$95.9\pm7.9\ ^{4}$	-
NA ratings, unpleasant IAPS images	$3.5\pm0.7~^5$	$3.6\pm0.6\ ^6$
NA ratings, neutral IAPS images	1.1 \pm 0.2 5	1.3 ± 0.3^{-6}
Time between MRI and CA-IMT assessments (absolute days)	23.7 ± 47.0	58.1 ± 35.0
CA-IMT (mm)	0.638 ± 0.117 8	0.609 ± 0.086 ⁸

Table 1. Analytic sample characteristics and descriptive statistics for study cohorts

Data are presented as percentage or mean ± standard deviation. AHAB-2, Adult Health and Behavior Project, Phase 2; PIP, Pittsburgh Imaging Project; HDL, highdensity lipoproteins; SBP, systolic blood pressure; DBP, diastolic blood pressure; NA, negative affect; CA-IMT, carotid-artery intima-media thickness. ¹N=457; ²N=454; ³N=424; ⁴N=451; ⁵N=183; ⁶N=158; ⁷N=155. ⁸In AHAB-2, 49 participants completed the CA-IMT assessment on the same day as MRI; 177 completed CA-IMT the assessment on a separate day prior to MRI; 232 completed the CA-IMT assessment on a separate day following MRI. In PIP, all participants completed the CA-IMT assessment on a separate day prior to MRI.